### PolyChord: Next Generation Nested Sampling Sampling, Parameter Estimation and Bayesian Model Comparison

Will Handley wh260@cam.ac.uk

Supervisors: Anthony Lasenby & Mike Hobson Astrophysics Department Cavendish Laboratory University of Cambridge

December 11, 2015

Parameter estimation & model comparison

Metropolis Hastings

Nested Sampling

PolyChord

Applications

► Data: D

- ► Data: D
- ► Model: *M*

- ► Data: D
- ► Model: *M*
- Parameters: Θ

- Data: D
- Model: M
- Parameters: Θ
- Likelihood:  $P(D|\Theta, M) = \mathcal{L}(\Theta)$

- Data: D
- Model: M
- Parameters: Θ
- Likelihood:  $P(D|\Theta, M) = \mathcal{L}(\Theta)$
- Posterior:  $P(\Theta|D, M) = \mathcal{P}(\Theta)$

- Data: D
- Model: M
- Parameters: Θ
- Likelihood:  $P(D|\Theta, M) = \mathcal{L}(\Theta)$
- Posterior:  $P(\Theta|D, M) = \mathcal{P}(\Theta)$
- Prior:  $P(\Theta|M) = \pi(\Theta)$

- Data: D
- Model: M
- Parameters: Θ
- Likelihood:  $P(D|\Theta, M) = \mathcal{L}(\Theta)$
- Posterior:  $P(\Theta|D, M) = \mathcal{P}(\Theta)$
- Prior:  $P(\Theta|M) = \pi(\Theta)$
- Evidence: P(D|M) = Z



Parameter estimation



Parameter estimation

What does the data tell us about the params  $\Theta$  of our model *M*?





$$\pi(\Theta) = \mathrm{P}(\Theta|M) \xrightarrow{D} \mathrm{P}(\Theta|D, M) = \mathcal{P}(\Theta)$$

$$\pi(\Theta) = \mathrm{P}(\Theta|M) \xrightarrow{D} \mathrm{P}(\Theta|D,M) = \mathcal{P}(\Theta)$$

Solution: Use the likelihood  $\mathcal{L}$  via Bayes' theorem:

$$\pi(\Theta) = \mathrm{P}(\Theta|M) \xrightarrow{D} \mathrm{P}(\Theta|D, M) = \mathcal{P}(\Theta)$$

Solution: Use the likelihood  $\mathcal{L}$  via Bayes' theorem:

$$P(\Theta|D, M) = \frac{P(D|\Theta, M)P(\Theta|M)}{P(D|M)}$$

$$\pi(\Theta) = \mathrm{P}(\Theta|M) \xrightarrow{D} \mathrm{P}(\Theta|D,M) = \mathcal{P}(\Theta)$$

Solution: Use the likelihood  $\mathcal{L}$  via Bayes' theorem:

$$P(\Theta|D, M) = \frac{P(D|\Theta, M)P(\Theta|M)}{P(D|M)}$$

$$\mathsf{Posterior} \ = \frac{\mathsf{Likelihood} \times \mathsf{Prior}}{\mathsf{Evidence}}$$

## Bayes' theorem

Model comparison



What does the data tell us about our model  $M_i$  in relation to other models  $\{M_1, M_2, \dots\}$ ?

### Bayes' theorem Model comparison

What does the data tell us about our model  $M_i$  in relation to other models  $\{M_1, M_2, \dots\}$ ?

 $\mathrm{P}(M_i) \xrightarrow{D} \mathrm{P}(M_i|D)$ 

### Bayes' theorem Model comparison

What does the data tell us about our model  $M_i$  in relation to other models  $\{M_1, M_2, \dots\}$ ?

 $\mathrm{P}(M_i) \xrightarrow{D} \mathrm{P}(M_i|D)$ 

$$\mathrm{P}(M_i|D) = \frac{\mathrm{P}(D|M_i)\mathrm{P}(M_i)}{\mathrm{P}(D)}$$

### Bayes' theorem Model comparison

What does the data tell us about our model  $M_i$  in relation to other models  $\{M_1, M_2, \dots\}$ ?

 $\mathrm{P}(M_i) \xrightarrow{D} \mathrm{P}(M_i|D)$ 

$$P(M_i|D) = \frac{P(D|M_i)P(M_i)}{P(D)}$$

 $P(D|M_i) = Z_i = Evidence of M_i$ 

# Parameter estimation & model comparison

The challenge

Parameter estimation: what does the data tell us about a model? (Computing posteriors)

Parameter estimation: what does the data tell us about a model? (Computing posteriors)

Model comparison: what does the data tell us about all models? (Computing evidences)

Parameter estimation: what does the data tell us about a model? (Computing posteriors) Model comparison: what does the data tell us about all models?

(Computing evidences)

Both of these are challenging things to compute.

Parameter estimation: what does the data tell us about a model? (Computing posteriors)

Model comparison: what does the data tell us about all models? (Computing evidences)

Both of these are challenging things to compute.

 Markov-Chain Monte-Carlo (MCMC) can solve the first of these (kind of)

Parameter estimation: what does the data tell us about a model? (Computing posteriors)

Model comparison: what does the data tell us about all models? (Computing evidences)

Both of these are challenging things to compute.

- Markov-Chain Monte-Carlo (MCMC) can solve the first of these (kind of)
- Nested sampling (NS) promises to solve both simultaneously.



1. In high dimensions, posterior  $\mathcal{P}$  occupies a vanishingly small region of the prior  $\pi$ .



- 1. In high dimensions, posterior  $\mathcal{P}$  occupies a vanishingly small region of the prior  $\pi$ .
- 2. Worse, you don't know where this region is.



- 1. In high dimensions, posterior  $\mathcal{P}$  occupies a vanishingly small region of the prior  $\pi$ .
- 2. Worse, you don't know where this region is.



Describing an N-dimensional posterior fully is impossible.

- 1. In high dimensions, posterior  $\mathcal{P}$  occupies a vanishingly small region of the prior  $\pi$ .
- 2. Worse, you don't know where this region is.



- Describing an N-dimensional posterior fully is impossible.
- Project/marginalise into 2- or 3-dimensions at best

- 1. In high dimensions, posterior  $\mathcal{P}$  occupies a vanishingly small region of the prior  $\pi$ .
- 2. Worse, you don't know where this region is.



- Describing an N-dimensional posterior fully is impossible.
- Project/marginalise into 2- or 3-dimensions at best
- Sampling the posterior is an excellent compression scheme.

# Markov-Chain Monte-Carlo (MCMC)

Metropolis-Hastings, Gibbs, Hamiltonian...

# Markov-Chain Monte-Carlo (MCMC)

Metropolis-Hastings, Gibbs, Hamiltonian...

► Turn the *N*-dimensional problem into a one-dimensional one.
- ► Turn the *N*-dimensional problem into a one-dimensional one.
- Explore the space via a biased random walk.

- ► Turn the *N*-dimensional problem into a one-dimensional one.
- Explore the space via a biased random walk.
  - 1. Pick random direction

- ► Turn the *N*-dimensional problem into a one-dimensional one.
- Explore the space via a biased random walk.
  - 1. Pick random direction
  - 2. Choose step length

- ► Turn the *N*-dimensional problem into a one-dimensional one.
- Explore the space via a biased random walk.
  - 1. Pick random direction
  - 2. Choose step length
  - 3. If uphill, make step...

- ► Turn the *N*-dimensional problem into a one-dimensional one.
- Explore the space via a biased random walk.
  - 1. Pick random direction
  - 2. Choose step length
  - 3. If uphill, make step...
  - 4. ... otherwise sometimes make step.

### MCMC in action



### MCMC in action



#### When MCMC fails Burn in



#### When MCMC fails Burn in



Tuning the proposal distribution



Tuning the proposal distribution



Multimodality



Multimodality



Phase transitions



The real reason...

The real reason...

The real reason...

$$\mathcal{Z} = \mathrm{P}(D|M)$$

The real reason...

$$\begin{aligned} \mathcal{Z} &= \mathrm{P}(D|M) \\ &= \int \mathrm{P}(D|\Theta, M) \mathrm{P}(\Theta|M) d\Theta \end{aligned}$$

The real reason...

$$\begin{aligned} \mathcal{Z} &= \mathrm{P}(D|M) \\ &= \int \mathrm{P}(D|\Theta, M) \mathrm{P}(\Theta|M) d\Theta \\ &= \int \mathcal{L}(\Theta) \pi(\Theta) d\Theta \end{aligned}$$

The real reason...

$$\begin{split} \mathcal{Z} &= \mathrm{P}(D|M) \\ &= \int \mathrm{P}(D|\Theta, M) \mathrm{P}(\Theta|M) d\Theta \\ &= \int \mathcal{L}(\Theta) \pi(\Theta) d\Theta \\ &= \langle \mathcal{L} \rangle_{\pi} \end{split}$$

The real reason...

MCMC does not give you evidences!

$$egin{aligned} \mathcal{Z} &= \mathrm{P}(D|M) \ &= \int \mathrm{P}(D|\Theta, M) \mathrm{P}(\Theta|M) d\Theta \ &= \int \mathcal{L}(\Theta) \pi(\Theta) d\Theta \ &= \langle \mathcal{L} 
angle_{\pi} \end{aligned}$$

 MCMC fundamentally explores the posterior, and cannot average over the prior.

John Skilling's alternative to MCMC!

John Skilling's alternative to MCMC!

New procedure:

John Skilling's alternative to MCMC!

New procedure: Maintain a set S of n samples, which are sequentially updated:

John Skilling's alternative to MCMC!

New procedure:

Maintain a set S of n samples, which are sequentially updated:

 $S_0$ : Generate *n* samples from the prior  $\pi$ .

John Skilling's alternative to MCMC!

New procedure:

Maintain a set S of n samples, which are sequentially updated:

 $S_0$ : Generate *n* samples from the prior  $\pi$ .

 $S_{n+1}$ : Delete the lowest likelihood sample in  $S_n$ , and replace it with a new sample with higher likelihood

John Skilling's alternative to MCMC!

New procedure:

Maintain a set S of n samples, which are sequentially updated:

 $S_0$ : Generate *n* samples from the prior  $\pi$ .

 $S_{n+1}$ : Delete the lowest likelihood sample in  $S_n$ , and replace it with a new sample with higher likelihood

Requires one to be able to sample from the prior, subject to a *hard likelihood constraint*.




















































































Why bother?

Why bother?

• At each iteration, the likelihood contour will shrink in volume by a factor of  $\approx 1/n$ .

Why bother?

- At each iteration, the likelihood contour will shrink in volume by a factor of  $\approx 1/n$ .
- Nested sampling zooms in to the peak of the posterior exponentially.

Why bother?

- At each iteration, the likelihood contour will shrink in volume by a factor of  $\approx 1/n$ .
- Nested sampling zooms in to the peak of the posterior exponentially.
- Nested sampling can be used to get evidences!

## Calculating evidences

## Calculating evidences

$$\mathcal{Z} = \int \mathcal{L}( heta) \pi( heta) d heta$$

#### Calculating evidences

Transform to 1 dimensional integral

$$\mathcal{Z} = \int \mathcal{L}( heta) \pi( heta) d heta$$
• Transform to 1 dimensional integral  $\pi(\theta)d\theta = dX$ 

$$\mathcal{Z} = \int \mathcal{L}( heta) \pi( heta) d heta$$

• Transform to 1 dimensional integral  $\pi(\theta)d\theta = dX$ 

$$\mathcal{Z} = \int \mathcal{L}(\theta) \pi(\theta) d\theta = \int \mathcal{L}(X) dX$$

• Transform to 1 dimensional integral  $\pi(\theta)d\theta = dX$ 

$$\mathcal{Z} = \int \mathcal{L}(\theta) \pi(\theta) d\theta = \int \mathcal{L}(X) dX$$

► X is the prior volume

• Transform to 1 dimensional integral  $\pi(\theta)d\theta = dX$ 

$$\mathcal{Z} = \int \mathcal{L}(\theta) \pi(\theta) d\theta = \int \mathcal{L}(X) dX$$

► X is the prior volume

$$X(\mathcal{L}) = \int_{\mathcal{L}( heta) > \mathcal{L}} \pi( heta) d heta$$

• Transform to 1 dimensional integral  $\pi(\theta)d\theta = dX$ 

$$\mathcal{Z} = \int \mathcal{L}(\theta) \pi(\theta) d\theta = \int \mathcal{L}(X) dX$$

X is the prior volume

$$X(\mathcal{L}) = \int_{\mathcal{L}( heta) > \mathcal{L}} \pi( heta) d heta$$

 i.e. the fraction of the prior which the iso-likelihood contour L encloses.































Evidence error

$$\Delta \log X \sim -\frac{1}{n}$$

Evidence error

$$\Delta \log X \sim -\frac{1}{n} \pm \frac{1}{n}$$

Evidence error

$$\Delta \log X \sim -\frac{1}{n} \pm \frac{1}{n}$$
  
 $\log X_i \sim -\frac{i}{n}$ 

Evidence error

$$\Delta \log X \sim -\frac{1}{n} \pm \frac{1}{n}$$
  
 $\log X_i \sim -\frac{i}{n} \pm \frac{\sqrt{i}}{n}$ 

Evidence error

$$\Delta \log X \sim -\frac{1}{n} \pm \frac{1}{n}$$
  
 $\log X_i \sim -\frac{i}{n} \pm \frac{\sqrt{i}}{n}$ 

• # of steps to get to H:  
$$i_H \sim nH$$

Evidence error

$$\Delta \log X \sim -\frac{1}{n} \pm \frac{1}{n}$$
  
 $\log X_i \sim -\frac{i}{n} \pm \frac{\sqrt{i}}{n}$ 

▶ # of steps to get to H:  
$$i_H \sim nH$$

► estimate of volume at *H*:  
$$\log X_H \approx -H \pm \sqrt{\frac{H}{n}}$$

Evidence error

approximate compression:

$$\Delta \log X \sim -\frac{1}{n} \pm \frac{1}{n}$$
  
 $\log X_i \sim -\frac{i}{n} \pm \frac{\sqrt{i}}{n}$ 

▶ # of steps to get to H:  
$$i_H \sim nH$$

• estimate of volume at *H*:  
$$\log X_H \approx -H \pm \sqrt{\frac{H}{n}}$$

estimate of evidence error:

$$\log \mathcal{Z} \approx \sum w_i \mathcal{L}_i \pm \sqrt{\frac{H}{n}}$$

Parameter estimation

Parameter estimation

NS can also be used to sample the posterior

Parameter estimation

- NS can also be used to sample the posterior
- The set of dead points are posterior samples with an appropriate weighting factor

#### When NS succeeds



#### When NS suceeds



## When NS succeeds



#### When NS suceeds



## Sampling from a hard likelihood constraint
#### Sampling from a hard likelihood constraint

"It is not the purpose of this introductory paper to develop the technology of navigation within such a volume. We merely note that exploring a hard-edged likelihood-constrained domain should prove to be neither more nor less demanding than exploring a likelihood-weighted space."

- John Skilling

## Sampling within an iso-likelihood contour

Previous attempts

Rejection Sampling MultiNest; F. Feroz & M. Hobson (2009).

Rejection Sampling MultiNest; F. Feroz & M. Hobson (2009).

Suffers in high dimensions

Rejection Sampling MultiNest; F. Feroz & M. Hobson (2009).
 Suffers in high dimensions
 Hamiltonian sampling F. Feroz & J. Skilling (2013).

Rejection Sampling MultiNest; F. Feroz & M. Hobson (2009).
 Suffers in high dimensions
 Hamiltonian sampling F. Feroz & J. Skilling (2013).
 Requires gradients and tuning

Rejection Sampling MultiNest; F. Feroz & M. Hobson (2009).
Suffers in high dimensions
Hamiltonian sampling F. Feroz & J. Skilling (2013).
Requires gradients and tuning
Diffusion Nested Sampling B. Brewer et al. (2009).

Rejection Sampling MultiNest; F. Feroz & M. Hobson (2009).
Suffers in high dimensions
Hamiltonian sampling F. Feroz & J. Skilling (2013).
Requires gradients and tuning
Diffusion Nested Sampling B. Brewer et al. (2009).
Very promising

Rejection Sampling MultiNest; F. Feroz & M. Hobson (2009).

Suffers in high dimensions

Hamiltonian sampling F. Feroz & J. Skilling (2013).

Requires gradients and tuning

Diffusion Nested Sampling B. Brewer et al. (2009).

- Very promising
- Too many tuning parameters


































































• This procedure satisfies detailed balance.

- > This procedure satisfies detailed balance.
- Works even if  $\mathcal{L}_0$  contour is disjoint.

- > This procedure satisfies detailed balance.
- Works even if  $\mathcal{L}_0$  contour is disjoint.

- This procedure satisfies detailed balance.
- Works even if  $\mathcal{L}_0$  contour is disjoint.
- ► Need N reasonably large ~ O(n<sub>dims</sub>) so that x<sub>N</sub> is de-correlated from x<sub>1</sub>.

Issues with Slice Sampling

## Issues with Slice Sampling

1. Does not deal well with correlated distributions.

## Issues with Slice Sampling

- 1. Does not deal well with correlated distributions.
- 2. Need to "tune" w parameter.



Correlated distributions

We make an affine transformation to remove degeneracies, and "whiten" the space.

- We make an affine transformation to remove degeneracies, and "whiten" the space.
- Samples remain uniformly sampled

- We make an affine transformation to remove degeneracies, and "whiten" the space.
- Samples remain uniformly sampled
- We use the covariance matrix of the live points and all inter-chain points

- We make an affine transformation to remove degeneracies, and "whiten" the space.
- Samples remain uniformly sampled
- We use the covariance matrix of the live points and all inter-chain points
- Cholesky decomposition is the required skew transformation

- We make an affine transformation to remove degeneracies, and "whiten" the space.
- Samples remain uniformly sampled
- We use the covariance matrix of the live points and all inter-chain points
- Cholesky decomposition is the required skew transformation
- w = 1 in this transformed space

Parallelised up to number of live points with openMPI.

- Parallelised up to number of live points with openMPI.
- Novel method for identifying and evolving modes separately.

- Parallelised up to number of live points with openMPI.
- Novel method for identifying and evolving modes separately.
- Implemented in CosmoMC, as "CosmoChord", with fast-slow parameters.











Primordial power spectrum  $\mathcal{P}_{\mathcal{R}}(k)$  reconstruction

#### ► Temperature data TT+lowP

- Temperature data TT+lowP
- ► Foreground (14) & cosmological (4 + 2 \* N<sub>knots</sub> 2) parameters

- Temperature data TT+lowP
- ► Foreground (14) & cosmological (4 + 2 \* N<sub>knots</sub> 2) parameters
- Marginalised plots of  $\mathcal{P}_{\mathcal{R}}(k)$

- Temperature data TT+lowP
- ► Foreground (14) & cosmological (4 + 2 \* N<sub>knots</sub> 2) parameters
- Marginalised plots of  $\mathcal{P}_{\mathcal{R}}(k)$

$$\mathrm{P}(\mathcal{P}_{\mathcal{R}}|k, N_{\mathrm{knots}}) = \int \delta(\mathcal{P}_{\mathcal{R}} - f(k; \theta)) \mathcal{P}(\theta) d\theta$$

#### 0 internal knots

Primordial power spectrum  $\mathcal{P}_{\mathcal{R}}(k)$  reconstruction



#### 1 internal knots

Primordial power spectrum  $\mathcal{P}_{\mathcal{R}}(k)$  reconstruction


Primordial power spectrum  $\mathcal{P}_{\mathcal{R}}(k)$  reconstruction



Primordial power spectrum  $\mathcal{P}_{\mathcal{R}}(k)$  reconstruction













### **Bayes Factors**



# Marginalised plot

Primordial power spectrum  $\mathcal{P}_{\mathcal{R}}(k)$  reconstruction



# Object detection

#### Toy problem



# Object detection

Evidences

#### Object detection Evidences

▶ 
$$\log Z$$
 ratio:  $-251 : -156 : -114 : -117 : -136$ 

#### Object detection Evidences

- ► log Z ratio: -251: -156: -114: -117: -136
- odds ratio:  $10^{-60}$  :  $10^{-19}$  : 1 : 0.04 :  $10^{-10}$

# PolyChord vs. MultiNest

Gaussian likelihood



# PolyChord vs. MultiNest

Gaussian likelihood



The future of nested sampling

We are at the beginning of a new era of sampling algorithms

- We are at the beginning of a new era of sampling algorithms
- Plenty of more work in to be done in exploring new versions of nested sampling

- We are at the beginning of a new era of sampling algorithms
- Plenty of more work in to be done in exploring new versions of nested sampling
- Nested sampling is just the beginning

- We are at the beginning of a new era of sampling algorithms
- Plenty of more work in to be done in exploring new versions of nested sampling
- Nested sampling is just the beginning
- arXiv:1506.00171

- We are at the beginning of a new era of sampling algorithms
- Plenty of more work in to be done in exploring new versions of nested sampling
- Nested sampling is just the beginning
- arXiv:1506.00171
- http://ccpforge.cse.rl.ac.uk/gf/project/polychord/