

Bayesian evidence of non-standard inflation: Isocurvature perturbations and running spectral index

Tommaso Giannantonio^{1,2,*} and Eiichiro Komatsu^{3,4}

¹Ludwig-Maximilians-Universität München, Universitäts-Sternwarte München, Scheinerstr. 1, D-81679 München, Germany

²Excellence Cluster Universe, Boltzmannstr. 2, D-85748 Garching, Germany

³Max-Planck-Institut für Astrophysik, Karl-Schwarzschild-Str. 1, D-85748 Garching, Germany

⁴Kavli Institute for the Physics and Mathematics of the Universe, Todai Institutes for Advanced Study, the University of Tokyo, Kashiwa, Japan 277-8583 (Kavli IPMU, WPI)

Bayesian model comparison penalizes models with more free parameters that are allowed to vary over a wide range, and thus offers the most robust method to decide whether some given data require new parameters. In this paper, we ask a simple question: do current cosmological data require extensions of the simplest single-field inflation models? Specifically, we calculate the Bayesian evidence of a totally anti-correlated isocurvature perturbation and a running spectral index of the scalar curvature perturbation. These parameters are motivated by recent claims that the observed temperature anisotropy of the cosmic microwave background on large angular scales is too low to be compatible with the simplest inflation models. Both a subdominant, anti-correlated cold dark matter isocurvature component and a negative running index succeed in lowering the large-scale temperature power spectrum. We show that the introduction of isocurvature perturbations is disfavored, whereas that of the running spectral index is only moderately favored, even when the BICEP2 data are included in the analysis without any foreground subtraction.

I. INTRODUCTION

Suppose that we wish to decide whether some data require the addition of a new parameter to a model. We may compare the logarithms of the likelihood values evaluated at the best-fit parameters. For example, the conventional χ^2 method uses $\Delta\chi^2 \equiv -2\ln(\mathcal{L}_1/\mathcal{L}_2)$. The obvious problem of this approach is that the addition of a new parameter is guaranteed to improve the fit, yielding a smaller χ^2 value. But then, what does $\Delta\chi^2$ mean when we find, say, $\Delta\chi^2 = -7$ by adding one more parameter? Do the data require such a parameter?

To address this issue, some criteria for comparing models have been discussed in the literature. The Akaike information criterion (AIC; [1]) and the Bayesian information criterion (BIC; [2]) penalize models with more parameters by adding to χ^2 a term proportional to the number of parameters. These criteria penalize all parameters equally regardless of predictability. For example, consider two parameters, one being allowed to vary from -1 to 1 , and the other from 0 to 10^{10} . While AIC and BIC penalize both parameters equally, a more sensible criterion should penalize the latter more strongly.

In this paper, we shall apply *Bayesian model comparison* [3] to test whether extensions of the simplest inflation models are required by the current cosmological data. The Bayesian model comparison penalizes models with more free parameters that are allowed to vary over a wide range. Specifically, we compute the Bayesian evidence, \mathcal{Z} , defined by

$$\mathcal{Z} \equiv \int d^N\theta \mathcal{L}(\text{data}|\theta)P(\theta), \quad (1)$$

where $\mathcal{L}(\text{data}|\theta)$ is the likelihood of the data given the model parameters θ , and $P(\theta)$ is the prior probability. We then compare two models by computing the logarithm of the ratio of their evidences, $\ln B \equiv \ln(\mathcal{Z}_1/\mathcal{Z}_2)$. Since the prior probability is normalized as $\int d^N\theta P(\theta) = 1$, $P(\theta)$ at a given set of θ becomes small when a model contains more parameters varying over a wide range. This gives that model a small \mathcal{Z} , hence penalizing it more strongly. The factor $\ln B$ can be interpreted as the mathematical odds between the models given the data, which can also be expressed heuristically using the so-called “Jeffrey’s scale”, according to which the evidence for (or against) a model is said to be weak, moderate, and strong if $\ln B > 1, 2.5,$ and $5,$ respectively [4]. We shall adopt Jeffrey’s scale throughout this paper.

Why consider extensions of the simplest inflation models? Here, the “simplest inflation models” refer to inflation models driven by a single scalar field with a simple potential yielding approximately a power-law power spectrum of the scalar curvature perturbation.

A detection of isocurvature modes of any form would rule out all single-field inflation models. Moreover, a detection of a cold dark matter (CDM) isocurvature mode would shed light on the nature of CDM, e.g., axions [5].

Given that the measured deviation of the scalar curvature power spectrum from scale invariance is $1 - n_s \simeq 0.04$ [6, 7], the running spectral index, $\rho_s \equiv dn_s/d\ln k$, is typically of order $(1 - n_s)^2 = \mathcal{O}(10^{-3})$; however, larger values are possible if the third derivative of the potential of a scalar field driving inflation is large [8]. Thus, a large running index of order 10^{-2} necessarily requires a new energy scale in the potential (or a kinetic term of the field [9]), making the models more complicated.

A motivation to consider these extensions of the simplest single-field inflation models comes from the observational data of the cosmic microwave background

* tommaso.giannantonio@usm.lmu.de

(CMB). The *Planck* collaboration claims that the CMB temperature power spectrum data that they obtain at low multipoles are too low to be compatible with the best-fit power-law ($\rho_s = 0$) adiabatic curvature perturbation spectrum [7]. Both a negative running index and a nearly scale-invariant CDM isocurvature component that is anti-correlated with the curvature perturbation can lower the low-multipole power, reducing this apparent “tension” in the *Planck* temperature data [10].

This tension is exacerbated [11], if a significant fraction of the B-mode polarization detected at degree angular scales by the BICEP2 collaboration [12] originates from the primordial, nearly scale-invariant gravitational waves generated during inflation, as such gravitational waves add extra power to the temperature power spectrum at low multipoles [13]. Then, do the *Planck* and BICEP2 data require either a negative running index or an anti-correlated CDM isocurvature perturbation? This is the question that we shall address in this paper using Bayesian model comparison.

Ref. [14] computed the Bayesian evidence of a running index, showing that evidence for running is insignificant. Our results differ from theirs because of the choice of the data set and the prior probability on the amplitude of gravitational waves. Ref. [15] computed $\Delta\chi^2$ for isocurvature perturbations, but did not perform a Bayesian model comparison; thus, they were unable to conclude whether the data required isocurvature perturbations.

The structure of this paper is as follows: We describe the models in Section II, and present the data sets we use and the analysis method in Section III. We describe our results in Section IV, and conclude in Section V.

II. MODELS

A. Model I: Running scalar spectral index

We write the scalar curvature power spectrum as

$$\mathcal{P}_{\mathcal{R}}(k) = A_s \bar{k}^{n_s - 1 + \frac{1}{2}\rho_s \ln \bar{k}}, \quad (2)$$

where n_s and ρ_s are the scalar spectral index and its running, respectively, and $\bar{k} \equiv k/(0.05 \text{ Mpc}^{-1})$ is the normalized wavenumber. The tensor power spectrum is

$$\mathcal{P}_h(k) = r_{0.05} A_s \bar{k}^{-r_{0.05}/8}, \quad (3)$$

where $r_{0.05}$ is the tensor-to-scalar ratio defined at $k = 0.05 \text{ Mpc}^{-1}$.

In the top panel of Fig. 1 we compare the temperature power spectrum data, $D_l \equiv l(l+1)C_l/(2\pi)$, measured by *Planck* [16] with three representative models. The solid line shows the best-fit six-parameter adiabatic Λ CDM model with $\rho_s = 0$ and $r_{0.05} = 0$. The short-dashed line is the sum of the solid line and the tensor temperature power spectrum with $r_{0.05} = 0.2$, showing how adding the tensor power spectrum with the tensor-to-scalar ratio suggested by the BICEP2 data (without

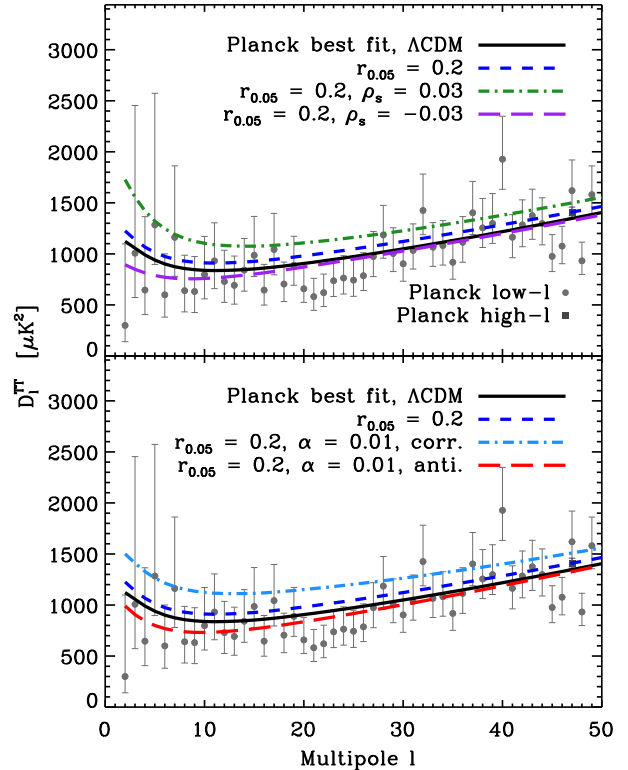


FIG. 1. Comparison of the Λ CDM and extended models. In both panels, the solid lines show the scalar CMB power spectrum of the six-parameter Λ CDM model, while the short-dashed lines show the sum of the solid lines and the tensor power spectrum with a tensor-to-scalar ratio of $r_{0.05} = 0.2$. The symbols with error bars show the *Planck* measurements [16]. (Top panel:) The long-dashed and dot-dashed lines show the sum of the tensor power spectrum and the scalar power spectrum with negative and positive running indices, respectively, with $|\rho_s| = 0.03$. (Bottom panel:) The long-dashed and dot-dashed lines show the sum of the short-dashed line and totally anti-correlated and correlated CDM isocurvature components, respectively, with an isocurvature-to-curvature ratio of $\alpha = 0.01$.

foreground subtraction) exacerbates the tension between the model and the *Planck* temperature data. The long-dashed line has $r_{0.05} = 0.2$ and a negative running index of $\rho_s = -0.03$, which brings the model back in agreement with the data. The dot-dashed line has a positive running index, yielding a bad fit.

B. Model II: CDM isocurvature

When we study an isocurvature component, we use Eq. (2) for the scalar curvature power spectrum with $\rho_s \equiv 0$. We continue to use the same tensor power spectrum as Eq. (3). We write the power spectrum of an isocurvature

Parameter	Description	Priors
$\omega_b \equiv \Omega_b h^2$	baryonic energy density	[0.020, 0.025]
$\omega_c \equiv \Omega_c h^2$	dark matter energy density	[0.080, 0.016]
100ϑ	sound horizon at last scattering	[1.34, 1.045]
τ	optical depth	[0.05, 0.18]
n_s	scalar spectral index	[0.90, 1.05]
$\log(10^{10} A_s)$	scalar amplitude	[2.9, 3.2]
$r_{0.05}$	tensor-to-scalar ratio	[0.0, 1.0]
α	isocurvature-to-curvature ratio	[0.0, 1.0]
ρ_s	scalar running spectral index	[-0.1, 0.1]

TABLE I. Parameters considered and prior ranges. In addition to these, all standard *Planck* nuisance parameters are left free and marginalized over.

component, \mathcal{S} , as

$$\mathcal{P}_{\mathcal{S}}(k) = \alpha A_s \bar{k}^{n_{\text{iso}}-1}, \quad (4)$$

where n_{iso} is the corresponding spectral index, and α is the isocurvature-to-curvature power ratio at $k = 0.05 \text{ Mpc}^{-1}$. We shall assume that \mathcal{R} and \mathcal{S} are totally anti-correlated (or correlated) throughout this paper. We thus write the cross-correlation power spectrum between \mathcal{R} and \mathcal{S} as

$$\mathcal{P}_{\mathcal{RS}}(k) = \pm \sqrt{\mathcal{P}_{\mathcal{R}}(k)\mathcal{P}_{\mathcal{S}}(k)}. \quad (5)$$

To minimize the number of parameters, we set $n_{\text{iso}} = n_s$.

In the lower panel of Fig. 1, the solid line shows the best-fit six-parameter adiabatic Λ CDM model with $\alpha = 0$ and $r_{0.05} = 0$. The short-dashed line is the sum of the black line and the tensor temperature power spectrum with $r_{0.05} = 0.2$, again showing that the BICEP2 data without foreground subtraction exacerbate the tension. The long-dashed line has $r_{0.05} = 0.2$ and a totally anti-correlated isocurvature component with $\alpha = 0.01$, which brings the model back in agreement with the data. The dot-dashed line has a totally correlated isocurvature component with $\alpha = 0.01$, yielding a bad fit.

III. DATA AND ANALYSIS METHOD

We use the *Planck* temperature power spectrum from the 2013 public release [16], with the addition of the WMAP 9-year polarization data [17] as combined in the default analysis by the *Planck* collaboration, as well as the B-mode polarization power spectrum released by the BICEP2 collaboration [12].

We also include a suite of baryon acoustic oscillation (BAO) distance scale measurements by the BOSS and 6dF collaborations, using the BOSS data release 9 (DR9) measurement at $z \simeq 0.57$ [18], the DR7 measurement at $z \simeq 0.35$ [19], and 6dF result at $z \simeq 0.1$ [20]. We do not use any supernovae or H_0 data.

We perform a Bayesian Monte Carlo exploration of the parameter space, using nested sampling as implemented in the public code MULTINEST [21, 22], used as

an alternative sampler within the COSMOMC/CAMB code [23, 24]. This method allows us to directly estimate the Bayesian evidence of each model and its uncertainties, and to compare them.

We let the parameters vary freely within the ranges described in Table I. As the nested sampling algorithm starts from uniform sampling over the whole parameter space, it is desirable to choose tight prior ranges such that the sampling is efficient. We thus choose a prior distribution for the standard Λ CDM parameters that is narrow, while being sufficiently broad so that the posterior likelihood of the six parameters is zero near the edges of the prior.

The prior distribution of the new parameters, i.e., $r_{0.05}$, α , and ρ_s , is chosen such that the power of tensor or isocurvature perturbations does not exceed that of the scalar curvature perturbation ($r_{0.05} \in [0, 1]$ and $\alpha \in [0, 1]$), and that the running spectral index is not too much bigger than $|1 - n_s|$ ($\rho_s \in [-0.1, 0.1]$). These prior distributions make physical sense and are compatible with expectations from inflation.

In addition to the parameters shown in Table I, we include the entire list of the standard *Planck* nuisance parameters, over which we marginalize. As in the standard *Planck* analysis, we account for massive neutrinos with a total mass fixed at $\sum m_\nu = 60 \text{ meV}$.

IV. RESULTS

A. Frequentist analysis: $\Delta\chi^2$

Let us first show the results from the frequentist analysis using the usual $\Delta\chi^2$ statistics. The sixth column of Table II shows $\Delta\chi^2$ values between Λ CDM+ $r_{0.05}$ and the other models. Negative values indicate a better fit over the former model. The first column shows the data combinations. When the BICEP2 data are included, we find $\Delta\chi^2 = -7.1$ and -4.2 for the running spectral index and the anti-correlated CDM isocurvature models, respectively. The isocurvature mode gives a smaller improvement because, while it reduces the low-multipole temperature power spectrum, it gives a non-negligible extra power near the first acoustic peak at $l \sim 200$, which is disfavored by the data.

Both models contain one more free parameter than Λ CDM+ $r_{0.05}$. While the $\Delta\chi^2$ values tell us that introducing one more parameter improves the fit, they do not tell us whether the data *require* such a parameter.

B. Bayesian evidences

Next, we show the results from the Bayesian analysis using the logarithms of the evidence ratio, $\ln B$. The seventh column of Table III shows $\ln B$ values between Λ CDM+ $r_{0.05}$ and the other models. *Positive* values indicate that the other models are favored over Λ CDM+ $r_{0.05}$.

Data	Model	Best fits	Best-fit χ^2	$\Delta\chi^2$ w.r.t. Λ CDM	$\Delta\chi^2$ w.r.t. $r\Lambda$ CDM
Planck + WP + BAO	Λ CDM	—	9804.1	—	0.0
	+ $r_{0.05}$	$r_{0.05} = 0.56 \cdot 10^{-3}$	9804.1	0.0	—
	+ α	$\alpha = 0.60 \cdot 10^{-3}$	9803.5	-0.6	-0.6
	+ ρ_s	$\rho_s = -0.0096$	9803.1	-1.0	-1.0
	+ $r_{0.05} + \alpha$	$r_{0.05} = 0.31 \cdot 10^{-3};$ $\alpha = 0.58 \cdot 10^{-3}$	9803.3	-0.8	-0.8
+ $r_{0.05} + \rho_s$	$r_{0.05} = 0.0014;$ $\rho_s = -0.012$	9802.8	-1.3	-1.3	
Planck + WP + BICEP2 B-mode + BAO	Λ CDM	—	9860.2	—	40.1
	+ $r_{0.05}$	$r_{0.05} = 0.16$	9820.1	-40.1	—
	+ α	$\alpha = 0.11 \cdot 10^{-2}$	9858.8	-1.4	38.7
	+ ρ_s	$\rho_s = -0.015$	9858.2	-2.0	38.1
	+ $r_{0.05} + \alpha$	$r_{0.05} = 0.17;$ $\alpha = 0.0024$	9815.9	-44.3	-4.2
+ $r_{0.05} + \rho_s$	$r_{0.05} = 0.18;$ $\rho_s = -0.026$	9813.0	-47.2	-7.1	

TABLE II. Frequentist analysis results.

Data	Model	95% c.l. posteriors	$\ln(Z)$	$\ln B = \Delta \ln Z$	Jeffrey's scale	$\ln B$ w.r.t. $r\Lambda$ CDM	Jeffrey's scale
Planck + WP + BAO	Λ CDM	—	-4940.57 \pm 0.04	—	—	2.82 \pm 0.04	moderate in favor
	+ $r_{0.05}$	$r_{0.05} \in [0, 0.12]$	-4943.39 \pm 0.02	-2.82 \pm 0.04	moderate against	—	—
	+ α	$\alpha \in [0, 0.0070]$	-4945.29 \pm 0.04	-4.72 \pm 0.06	moderate against	-1.90 \pm 0.04	weak against
	+ ρ_s	$\rho_s \in [-0.031, 0.0029]$	-4942.26 \pm 0.04	-1.69 \pm 0.06	weak against	1.13 \pm 0.04	weak in favor
	+ $r_{0.05} + \alpha$	$r_{0.05} \in [0, 0.19];$ $\alpha \in [0, 0.0097]$	-4947.45 \pm 0.02	-6.88 \pm 0.04	strong against	-4.06 \pm 0.03	moderate against
+ $r_{0.05} + \rho_s$	$r_{0.05} \in [0, 0.24];$ $\rho_s \in [-0.044, -0.0012]$	-4943.27 \pm 0.03	-2.70 \pm 0.05	moderate against	0.12 \pm 0.04	inconclusive	
Planck + WP + BICEP2 B-mode + BAO	Λ CDM	—	-4968.69 \pm 0.05	—	—	-17.26 \pm 0.06	strong against
	+ $r_{0.05}$	$r_{0.05} \in [0.096, 0.23]$	-4951.43 \pm 0.03	17.26 \pm 0.06	strong in favor	—	—
	+ α	$\alpha \in [0, 0.0078]$	-4973.09 \pm 0.14	-4.40 \pm 0.15	moderate against	-21.66 \pm 0.14	strong against
	+ ρ_s	$\rho_s \in [-0.033, 0.00063]$	-4969.86 \pm 0.03	-1.17 \pm 0.06	moderate against	-18.43 \pm 0.04	strong against
	+ $r_{0.05} + \alpha$	$r_{0.05} \in [0.12, 0.25];$ $\alpha \in [0, 0.014]$	-4953.52 \pm 0.06	15.17 \pm 0.08	strong in favor	-2.09 \pm 0.07	weak against
+ $r_{0.05} + \rho_s$	$r_{0.05} \in [0.11, 0.27];$ $\rho_s \in [-0.049, 0.011]$	-4948.88 \pm 0.03	19.81 \pm 0.06	strong in favor	2.55 \pm 0.04	moderate in favor	

TABLE III. Bayesian analysis results.

When the BICEP2 data are included, we find $\ln B = 2.55$ and -2.09 for the running spectral index and the CDM isocurvature models, respectively. These results clearly show the power of Bayesian model comparison: despite an improved χ^2 , the anti-correlated CDM isocurvature model is *disfavored* by the data. The running spectral index model is still favored, and it is “moderately favored” according to Jeffrey’s scale.

We show the marginalized 2D posteriors on the parameters of interest in Fig. 2, where we can see a visual confirmation of the 95% confidence intervals shown in the third column of Table III: the scalar running is favored at the 2σ level, while the amount of anti-correlated CDM isocurvature is consistent with zero.

We have tested the stability of our results when including the *Planck* CMB lensing likelihood, removing BAOs, and using $\mathcal{P}_h \propto k^0$ instead of $k^{-r/8}$. We find that the results are relatively robust, although the evidence in favor of running is reduced in some of these cases: the addition of CMB lensing in particular reduces the evidence to $\ln B = 1.6$, which is “weak” on Jeffrey’s scale. The *Planck* collaboration also finds a reduced significance of a running index when using the CMB lensing data [7].

Our results change more significantly if the same method of Ref. [14] is used, where the posterior likeli-

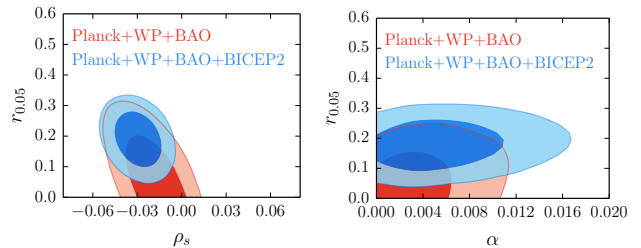


FIG. 2. Marginalized 2D posteriors on the tensor-scalar ratio, running, and isocurvature parameters.

hood of the tensor-to-scalar ratio obtained by the BICEP2 collaboration was used as a prior instead of calculating the full BICEP2 likelihood for each model. If we use their method, we reproduce their results, which show an even smaller evidence ratio for the running spectral index model, $\ln B = 1.1$. While applying the BICEP2 posterior distribution on $r_{0.05}$ as a prior is reasonable when constraining the tensor amplitude only, the results will be only approximately recovered if both $r_{0.05}$ and ρ_s are varied simultaneously. We thus conclude that Ref. [14] underestimated the evidence ratio for the running spectral index model.

V. CONCLUSIONS

There are at least three easy ways to reduce the apparent “tension” between the simplest inflation models with a tensor mode and the current CMB data including *Planck* and BICEP2. First, a sub-dominant CDM isocurvature perturbation anti-correlated with the dominant curvature perturbation [15]; second, a negative running spectral index [12]; and third, a modification of the large-scale primordial power spectrum [14, 25, 26].

We have performed a Bayesian model comparison of the former two extensions against the simplest inflation models. The anti-correlated CDM isocurvature component reduces the CMB temperature power spectrum at low multipoles, improving the agreement with the tensor model with $r_{0.05} = 0.2$ suggested by the BICEP2 data without any foreground subtraction. Nonetheless, we have found that such an improvement is Bayesianly *disfavored*, i.e., the data do not support such an extension of the inflation model, despite that it gives an improved χ^2 by $\Delta\chi^2 = -4.2$. This shows the power of the Bayesian model comparison method. While this result necessarily depends on the chosen prior on the amount of isocurvature, i.e., $\alpha \in [0, 1]$, this prior is physically motivated, and there is little room for ambiguity on the prior choice.

We have then tested a model with a running spectral index, as a negative running can also reduce the temperature power spectrum at low multipoles. We have found that a negative running spectral index is moderately favored with the log evidence ratio of $\ln B = 2.55$.

Our results are derived assuming that there is no foreground contamination in the BICEP2 data. Any foreground contributions will lower $\ln B$, and thus the anti-correlated CDM isocurvature will be even more disfavored, and the evidence for a negative running spectral index will likely turn to be “weak” ($\ln B < 2.5$). The BICEP2 collaboration finds that the polarized dust emission could account for 30% of the measured B-mode power spectrum, while others argue that 100% could be accounted for by dust [27, 28]. Therefore, we conclude that the current data do not require these particular extensions of the simplest inflation models.

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