THE THREE-POINT CORRELATION FUNCTION OF GALAXIES DETERMINED FROM THE TWO-DEGREE FIELD GALAXY REDSHIFT SURVEY

Y. P. Jing^{1, 2} and G. Börner² Received 2003 November 15; accepted 2004 February 4

ABSTRACT

In a detailed analysis of the three-point correlation function (3PCF) for the 2dF Galaxy Redshift Survey (2dFGRS), we have accurately measured the 3PCF for galaxies of different luminosities. The 3PCF amplitudes $[Q_{\text{red}}(s, u, v) \text{ or } Q_{\text{proj}}(r_p, u, v)]$ of the galaxies generally decrease with increasing triangle size and increase with the shape parameter v, in qualitative agreement with the predictions for the clustering of dark matter (DM) in popular hierarchical cold dark matter (CDM) models. The 2dFGRS results agree well with our results for the Las Campanas Redshift Survey (LCRS), although the measurement accuracy is greatly improved in the present study because the 2dFGRS is much larger in size than the LCRS. The dependence of the 3PCF on luminosity is not significant, but there seems to be a trend for the brightest galaxy sample to have a lower amplitude than the fainter ones. Comparing the measured 3PCF amplitudes $[Q_{red}(s, u, v)]$ or $Q_{proj}(r_p, u, v)$ to the prediction of a WMAP (Wilkinson Microwave Anisotropy Probe) concordance model, we find that the measured values are consistently lower than the predicted ones for DM. This is most pronounced for the brightest galaxies (sample I), for which about one-half of the predicted Q-value provides a good description of $Q_{\text{proj}}(r_p, u, v)$ for the 2dFGRS data. For the less luminous sample (sample II), the Q-values are also smaller than in the DM model on small scales, but on scales larger than $s = 8 \ h^{-1}$ Mpc and $r_p = 3.25 \ h^{-1}$ Mpc they reach the model values. Therefore, the galaxies of sample II are unbiased tracers on linear scales, but the bright galaxies (sample I) have a linear bias factor of ∼1.5. As for the LCRS data, we may state that the best-fit DM model gives higher values for the 3PCF than observed. This indicates that the simple DM models must be refined, by using either more sophisticated bias models or a more sophisticated combination of model parameters.

Subject headings: cosmology: theory — dark matter — galaxies: clusters: general — galaxies: distances and redshifts — large-scale structure of universe

On-line material: color figures

1. INTRODUCTION

To infer the spatial distribution of cosmic matter from the observed distribution of galaxies is a nontrivial task. Big redshift catalogs of galaxies and numerical simulations of the dark matter (DM) clustering, depending on the cosmological model and initial conditions, are the observational and theoretical bases for a treatment of this problem. The statistical properties of both the theoretical models and the observational catalogs can be obtained by some powerful tool, such as the *n*-point correlation functions (Peebles 1980, hereafter P80). The present state of the universe is thought to have evolved from initial conditions for the density field that are one specific realization of a random process, with the density contrast as the random variable. A Gaussian distribution for the initial conditions, such as is predicted by the inflationary scenario, is fully determined by the two-point correlation function (2PCF) or its Fourier transform, the power spectrum P(k).

This connection has motivated an extensive use of the 2PCF to analyze galaxy catalogs (e.g., Davis & Peebles 1983; Jing, Mo, & Börner 1998, hereafter JMB98; Hamilton & Tegmark 2002; Zehavi et al. 2002; Norberg et al. 2002a; Hawkins et al. 2003), the cosmic microwave background anisotropy (e.g., Spergel et al. 2003 and references therein), and the cosmic

shear field (e.g., Pen et al. 2003; Bartelmann & Schneider 2001 and references therein). Several constraints on theoretical models have already been derived, despite the fact that there are many ingredients to a specific model that can be optimally adapted to the properties of a given galaxy sample. The cosmological parameters, the initial power spectrum of the DM component, and the bias, i.e., the difference in the clustering of galaxies and DM particles, can all be adjusted to some extent.

The three-point correlation function (3PCF) $\zeta(r_{12}, r_{23}, r_{31})$ characterizes the clustering of galaxies in further detail (P80) and can provide additional constraints for cosmogonic models. The 3PCF is zero for a Gaussian field, but during the time evolution of the density perturbations the distribution develops non-Gaussian properties. These can be measured by the 3PCF or, equivalently, its Fourier-transformed counterpart, the bispectrum, and thus additional information on the nature of gravity and DM is gained, including an additional test of the structure formation models.

The theories based on cold dark matter (CDM) models predict that the 3PCF of galaxies depends on the shape of the linear power spectrum (Fry 1984; Jing & Börner 1997; Scoccimarro et al. 1998; Buchalter & Kamionkowski 1999) and the galaxy biasing relative to the underlying mass (Davis et al. 1985; Gaztañaga & Frieman 1994; Mo, Jing, & White 1997; Matarrese, Verde, & Heavens 1997; Catelan et al. 1998). The second-order perturbation theory (PT2) predicts that the 3PCF of the DM depends on the shape of the triangle formed by the three galaxies and on the slope of the linear

¹ Shanghai Astronomical Observatory, the Partner Group of MPI für Astrophysik, Nandan Road 80, Shanghai 200030, China.

Max-Planck-Institut f
ür Astrophysik, Karl-Schwarzschild-Strasse 1, 85748 Garching, Germany.

power spectrum (Fry 1984; Jing & Börner 1997; Barriga & Gaztañaga 2002; see Bernardeau et al. 2002 for an excellent review).

The determination of the 3PCF was pioneered by Peebles and his coworkers in the 1970s. They proposed a so-called hierarchical form

$$\zeta(r_{12}, r_{23}, r_{31}) = Q[\xi(r_{12})\xi(r_{23}) + \xi(r_{23})\xi(r_{31}) + \xi(r_{31})\xi(r_{12})],$$
(1)

with the constant $Q \approx 1.29 \pm 0.2$. This form is valid for scales of $r \lesssim 3 \ h^{-1}$ Mpc (P80). Subsequently, the analysis of several galaxy catalogs has supported this result. The ESO-Uppsala catalog of galaxies (Lauberts 1982) was analyzed by Jing, Mo, & Börner (1991). The 3PCF was also examined for the CfA, Anglo-Australian Telescope, and Koester-Oemler-Schechter-Shectman Survey redshift samples of galaxies (Peebles 1981; Bean et al. 1983; Efstathiou & Jedrzejewski 1984; Hale-Sutton et al. 1989). These earlier redshift samples are too small, with ≤ 2000 galaxies, to allow a test of the hierarchical form in redshift space. Only fits to the hierarchical form were possible. The Q-value obtained in this way from redshift samples is around 0.6 (Efstathiou & Jedrzejewski 1984), much smaller than the value extracted by Peebles and his coworkers from the Lick and Zwicky catalogs. Redshift distortion effects are probably responsible for this reduction (Matsubara 1994).

If the density field of the galaxies $\delta_g(\mathbf{x})$ is connected to the matter overdensity $\delta_m(\mathbf{x})$ as

$$\delta_g = b_1 \delta_m + b_2 \delta_m^2, \tag{2}$$

then in PT2 $P_g(k) = b_1^2 P_m(k)$, and

$$Q_a = Q_m/b_1 + b_2/b_1^2 (3)$$

for the Q-value of the galaxy 3PCF. Since Q_m depends on the shape of the power spectrum in PT2 that can be measured from the galaxy power spectrum on large scales (assuming a linear bias), one may measure the bias parameters b_1 and b_2 from the 3PCF of galaxies on large scales.

The hierarchical form (eq. [1]) is purely empirical without a solid theoretical argument supporting it. In contrast, the PT2 theory predicts that Q_m of DM depends on the shape of triangles on the linear clustering scale. Even in the strongly nonlinear regime where the hierarchical form was expected to hold, the CDM models do not seem to obey it, as demonstrated by Jing & Börner (1998, hereafter JB98). The large sample size of the Las Campanas Redshift Survey (LCRS; Shectman et al. 1996) made it possible for the first time to study the detailed dependence of the amplitude Q_q of galaxies on the shape and size of triangles. JB98 computed the 3PCFs for the LCRS in both redshift and projected space. As demonstrated by JB98, the projected 3PCF they proposed has simple relations to the real-space 3PCF. Their results have revealed that in both redshift and real space there are small, but significant, deviations from the hierarchical form.

The general dependence of the galaxy 3PCF on triangle shape and size appeared to be in qualitative agreement with the CDM cosmogonic models. JB98 found that a CDM model with $\Omega_m h = 0.2$, and an appropriately chosen bias scheme (the cluster-weighted model originally proposed in JMB98, now generally called the halo occupation number model in the literature) meets the constraints imposed by the LCRS data on the 2PCF and the pairwise velocity dispersion (PVD) of the

galaxies. The real-space Q_g obtained from the LCRS is, however, well described by half the mean Q_m -value predicted by this best-fit CDM model. The unavoidable conclusion is that it is difficult to find a simple model that meets all the constraints.

In recent years, several authors have measured the 3PCF and the bispectrum, with emphasis on the quasi-linear and linear clustering scales. For example, for the APM (Automated Plate-Measuring Machine) galaxies (Gaztañaga & Frieman 1994; Frieman & Gaztañaga 1999), the *IRAS* galaxies (Scoccimarro et al. 2001), and the 2dFGRS (Two-Degree Field Galaxy Redshift Survey) galaxies (Verde et al. 2002), the measurements were used to constrain the linear and nonlinear bias parameters b_1 and b_2 (eq. [2]), by comparison with a model for the 3PCF obtained in PT2.

For the APM galaxies the PT2 model for the 3PCF agrees well with the APM catalog measurements on large scales (Frieman & Gaztañaga 1999), which implies $b_1 \approx 1$ and $b_2 \approx 0$. The bispectrum of *IRAS* PSCz galaxies leads to values of

$$b_1^{-1} = 1.32(+0.36, -0.39),$$
 (4)

$$b_2^{-1} = 1.15 \pm 0.39 \tag{5}$$

(Scoccimarro et al. 2001) for the wavenumber k in the interval 0.05 h Mpc⁻¹ $\leq k \leq$ 0.2 h Mpc⁻¹. The measurement of the bispectrum for the 2dFGRS catalog resulted in bias parameters

$$b_1 = 1.04 \pm 0.11,$$
 (6)

$$b_2 = -0.054 \pm 0.08 \tag{7}$$

on scales between 5 and 30 h^{-1} Mpc (Verde et al. 2002). These results indicate that on large scales, optical galaxies (both 2dFGRS and APM galaxies) are unbiased relative to the underlying mass distribution, while the *IRAS* galaxies are an antibiased tracer. Furthermore, the nonlinear bias of the *IRAS* galaxies is significantly nonzero. Combining these results with our result on the LCRS (JB98) implies that optical galaxies are a biased tracer on small scales but an unbiased tracer on larger scales

In this paper, we measure the 3PCF in both redshift and projected space for the 2dFGRS (Colless et al. 2001, hereafter C01). We are motivated to investigate further the mismatch of the 3PCFs found by JB98 between the concordance CDM model and the LCRS. Because the 2dFGRS covers a much larger volume than the LCRS, we expect to measure the 3PCF more accurately, especially on large scales. Therefore, we attempt to find out if there exists a transition at which the 3PCF gradually approaches the unbiased prediction of the concordance CDM model on large scales (Frieman & Gaztañaga 1999; Verde et al. 2002) from half of the CDM prediction on small scales (JB98). The results of Frieman & Gaztañaga (1999) and Verde et al. (2002) apparently imply a high normalization, $\sigma_8 \approx 1$, for the primordial fluctuation (σ_8 is the linear rms density fluctuation at the present in a sphere of radius 8 h^{-1} Mpc), while some observations, e.g., the PVD of galaxies and the abundance of clusters of galaxies, clearly prefer a smaller value of $\sigma_8 \approx 0.7$ for the concordance ΛCDM model (e.g., Bahcall & Comerford 2002; Lahav et al. 2002; van den Bosch, Mo, & Yang 2003; Yang et al. 2004). This apparent conflict also motivates us to examine the 3PCF more carefully on quasi-linear scales, which can be explored by the 2dFGRS. Moreover, it is well known that the clustering of

TABLE 1								
SAMPLES	SELECTED	According	то	LUMINOSITY				

Sample	$M_b - 5 \log h$	South (no. of galaxies)	North (no. of galaxies)	Total (no. of galaxies)
I	≤−19.66	16702	11761	28,463
II	$-19.66 < M_b - 5 \log h \le -18.5$	14247	11798	26,045
III	≤−18.5	30949	23559	54,508
IV	>-18.5	7930	6572	14,502
V	No limit	39208	30447	69,655

galaxies depends on their luminosity. In Frieman & Gaztañaga (1999) and Verde et al. (2002), galaxies are included in a wide range of luminosity, and it is difficult to determine whether for some luminosity range galaxies are unbiased relative to the mass distribution on large scales. In this paper, we attempt to measure the 3PCF for the first time as a function of luminosity. We believe that these measurements of the 3PCF will provide useful observational constraints on galaxy formation theories.

In \S 2, we describe the sample selection for the analysis, the selection effects, and the procedure for generating random and mock samples. The statistical methods for measuring the 3PCFs are presented in \S 3. The results of the 2dFGRS are given in \S 4, along with a comparison with the results of the LCRS and the predictions of the concordance model for DM. Our results are summarized in \S 5.

2. OBSERVATIONAL SAMPLE, RANDOM SAMPLE, AND MOCK CATALOGS

We select data for our analysis from the 100K public release³ of the 2dFGRS (C01). The survey covers two declination strips, one in the south Galactic pole (SGP) and other in the north Galactic pole (NGP), and 99 random fields in the southern Galactic cap. In this paper, only galaxies in the two strips are considered. Further criteria for the inclusion of galaxies in our analysis are that they are within the redshift range of 0.02 < z < 0.25, have the redshift measurement quality $Q \ge 3$, and are in regions with redshift sampling completeness $R(\theta)$ better than 0.1 (where θ is sky position). The redshift range restriction is imposed so that the clustering statistics are less affected by the galaxies in the Local Supercluster and by the sparse sampling at high redshift. The redshift quality restriction is imposed so that only galaxies with reliable redshifts are used in our analysis. An additional reason is that the redshift completeness mask provided by the survey team, which is used in our analysis, is constructed for the redshift catalog of $Q \ge 3$. The last restriction is imposed in order to eliminate galaxies in the fields for which the field redshift completeness c_F is less than 70% [see C01 about the difference between $R(\theta)$ and c_F]. These fields are (or will be) reobserved and have not been included in computing the redshift mask map $R(\theta)$. Finally, there are a total of 69,655 galaxies satisfying our selection criteria, 30,447 in the NGP strip and 39,208 in the SGP strip.

It is well known that the two-point clustering of galaxies depends on the luminosity (Xia, Deng, & Zhou 1987; Börner et al. 1991; Loveday et al. 1995; Norberg et al. 2002a), and the luminosity dependence is an important constraint on galaxy

formation models (Kauffmann, Nusser, & Steinmetz 1997; Kauffmann et al. 1999; Benson et al. 2000; Yang, Mo, & van den Bosch 2003). We take advantage of the size of the 2dFGRS to carry out a first study of the luminosity dependence of the 3PCF. The galaxies are divided into three classes: luminous galaxies with absolute magnitude $M_b \leq M_b^* =$ $-19.66 + 5 \log h$, faint galaxies with $M_b > -18.5 + 5 \log h$, and typical galaxies with luminosity in between, where M_h^* is the characteristic luminosity of the Schechter function in the $b_{\rm J}$ band (Norberg et al. 2002b) and h is the Hubble constant in units of 100 km s⁻¹ Mpc⁻¹. We also do the analysis for galaxies with $M_b \le -18.5 + 5 \log h$ in order to compare the results with those of the previous study of the LCRS (JB98). The details of the subsamples studied in this paper are given in Table 1. For computing the absolute magnitude, we have used the k-correction and luminosity evolution model of Norberg et al. (2002b; k+e model), i.e., the absolute magnitude is in the rest-frame $b_{\rm J}$ band at z=0. We assume a cosmological model with the density parameter $\Omega_0 = 0.3$ and the cosmological constant $\lambda_0 = 0.7$ throughout this paper.

A detailed account of the observational selection effects has been released with the catalog by the survey team (C01). The limiting magnitude changes slightly across the survey region because of further magnitude calibrations that were carried out after the target galaxies had been selected for the redshift measurement. This observational effect is documented in the magnitude limit mask $b_{\rm J}^{\rm lim}(\pmb{\theta})$ (C01). The redshift sampling is far from uniform within the survey region, and this selection effect is given by the redshift completeness mask $R(\pmb{\theta})$. The redshift measurement success rate also depends on the brightness of galaxies, making fainter galaxies more incomplete in the redshift measurement. The $\mu(\pmb{\theta})$ mask provided by the survey team has the aim of accounting for the brightness-dependent incompleteness.

These observational effects can be corrected in our analysis for the 3PCF through properly generating random samples. To construct the random samples, we first select a spatial volume that is sufficiently large to contain the survey sample. Then we randomly distribute points within the volume and eliminate the points that are outside the survey boundary. Adopting $15.0 < b_J \le b_J^{\text{lim}}(\theta)$ for the magnitude limits⁴ of the survey in the direction θ , we select random points, according to the luminosity function of the 2dFGRS and the k+e model for the k-correction and luminosity evolution (Norberg et al. 2002a), and assign to each point an apparent magnitude (and an absolute magnitude). This unclustered sample is a random sample for the 2dFGRS photometric catalog. Then we

³ Available at http://www.mso.anu.edu.au/2dFGRS.

⁴ We assume that the brighter magnitude limit for the survey is 15.0. This is a reasonable value for the survey, but our results are insensitive to the choice of this value.

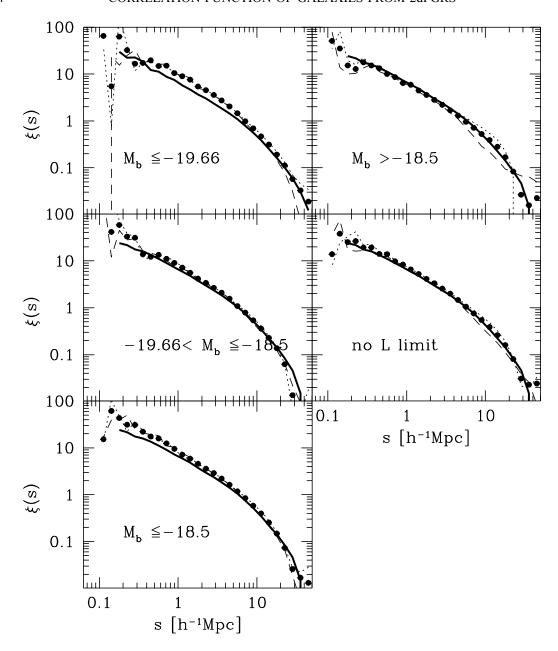


Fig. 1.—Redshift 2PCF of galaxies with different luminosities measured from the 2dFGRS catalog. The symbols with error bars are for the whole catalog, the dotted lines are for the SGP subsample, and the dashed lines are for the NGP subsample. The errors are estimated by the bootstrap resampling method. The luminosity ranges are indicated in each panel. The thick solid lines are for the simulation predictions for the redshift 2PCF of DM in the WMAP concordance CDM model at redshift z = 0.13. [See the electronic edition of the Journal for a color version of this figure.]

implement the magnitude-dependent redshift selection effect according to C01. We keep random points of magnitude m in the direction θ at a sampling rate $S(\theta, m)$, which reads as (eq. [11] of C01)

$$S(\boldsymbol{\theta}, m) = \frac{N_p(\boldsymbol{\theta}, m)}{N_e(\boldsymbol{\theta}, m)} R(\boldsymbol{\theta}, m) c_z[m, \mu(\boldsymbol{\theta}, m)], \qquad (8)$$

where $N_p(\theta, m)$ is the number of parent catalog galaxies in the sector θ and $N_e(\theta, m)$ is the number of galaxies that are expected to have measured redshifts for given $\mu(\theta, m)$. The ratio $N_e(\theta, m)/N_p(\theta, m)$ is actually the field completeness $c_F(\theta, m)$ defined by C01, which we compute according to their equation (7) (see also Norberg et al. 2002b). The function $R(\theta, m)$ is given by the redshift completeness mask, and $c_z[m, \mu(\theta, m)]$ can be easily computed from the μ mask

(eq. [5] of C01). We have used the corrected value 0.5 ln 10 for the α parameter in the power-law galaxy count model, according to the 2dFGRS Web page.⁵

We have checked the random samples carefully by reproducing the angular distribution, the mean redshift distribution, and especially the two-point statistics of clustering of the observed catalog. It is known that the 2PCF measured from galaxy catalogs on large scales is sensitive to the details of corrections for the above selection effects. We have estimated the redshift and projected correlation functions by the same method as in JMB98 for the LCRS. The 2PCFs are shown in Figures 1 and 2 and can be compared with the results of the 2dFGRS team for the clustering of galaxies (e.g., Hawkins

⁵ See http://www.mso.anu.edu.au/2dFGRS.

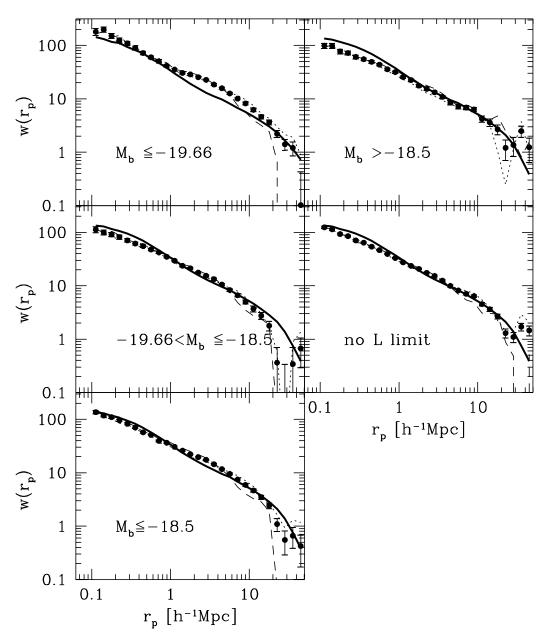


Fig. 2.—Same as Fig. 1, but for the projected 2PCF. [See the electronic edition of the Journal for a color version of this figure.]

et al. 2003; Norberg et al. 2002a). In addition to the broad agreement with their results, even the subtle difference between the northern and southern caps (the clustering on large scales is slightly larger in the southern cap than in the northern cap) and the luminosity dependence of the clustering are well reproduced in our analysis.

We did not take into account in our analysis the fiber collision effect that two galaxies closer than $\sim 30''$ cannot be assigned fibers simultaneously in one spectroscopic observation. Thus, one of them will not have a redshift observation if no reobservation is arranged. This effect reduces the real-space (or projected) 2PCF at small separations. With LCRS, JMB98 estimated the effect to lead to a 15% reduction in the 2PCF at projected separations of 100 h^{-1} kpc and to a less than 5% reduction at separations larger than 400 h^{-1} kpc. This effect is smaller (10% reduction at separations of 100 h^{-1} kpc) in the 2dFGRS (Hawkins et al. 2003), because the limiting fiber

separation is slightly smaller (30" in the 2dFGRS vs. 55" in the LCRS), and one field may be observed more than once in the 2dFGRS observation strategy. JB98 have examined the fiber collision effect on their measurement of the 3PCF of the LCRS. They found that the effect reduces the real-space (projected) 3PCF at small separations but changes little the normalized 3PCFs, Q, that we measure in this paper, because the effects on the 2PCF and 3PCF are canceled out when Q is measured. Since the effect is slightly smaller in 2dFGRS in terms of the two-point clustering, we believe that only a negligible effect on our measurement of the normalized three-point correlations would result.

3. STATISTICAL METHODS

We measure the 3PCFs for the galaxies in the 2dFGRS following the method of JB98. By definition, the joint probability dP_{123} of finding one object simultaneously in each of

the three volume elements $d\mathbf{r}_1$, $d\mathbf{r}_2$ and $d\mathbf{r}_3$, at positions \mathbf{r}_1 , \mathbf{r}_2 and \mathbf{r}_3 , respectively, is as follows (P80):

$$dP_{123} = \bar{n}(\mathbf{r}_1)\bar{n}(\mathbf{r}_2)\bar{n}(\mathbf{r}_3)\left[1 + \xi(r_{12}) + \xi(r_{23}) + \xi(r_{31}) + \zeta(r_{12}, r_{23}, r_{31})\right]d\mathbf{r}_1 d\mathbf{r}_2 d\mathbf{r}_3,$$
(9)

where $r_{ij} = |\mathbf{r}_i - \mathbf{r}_j|$, $\bar{n}(\mathbf{r}_i)$ is the mean density of galaxies at \mathbf{r}_i , and $\zeta(r_{12}, r_{23}, r_{31})$ is the 3PCF. This definition can be applied straightforwardly to redshift surveys of galaxies to measure the 3PCF $\zeta(s_{12}, s_{23}, s_{31})$ of galaxies in redshift space (at this point we neglect the anisotropy induced by the redshift distortion, which is considered below). Here and below we use \mathbf{r} to denote the real space and \mathbf{s} the redshift space.

The 3PCF of galaxies can be measured from the counts of different triplets (P80). Four types of distinct triplets, with triangles in the range $[s_{12} \pm (1/2)\Delta s_{12}, s_{23} \pm (1/2)\Delta s_{23},$ and $s_{31} \pm (1/2)\Delta s_{31}]$, are counted: $DDD(s_{12}, s_{23}, s_{31})$ for triplets formed by three galaxies; $DDR(s_{12}, s_{23}, s_{31})$ for triplets formed by two galaxies and one random point; $DRR(s_{12}, s_{23}, s_{31})$ for triplets formed by one galaxy and two random points; and $RRR(s_{12}, s_{23}, s_{31})$ for triplets formed by three random points. The random sample of points is generated in the way described in § 2. Following the definition in equation (9), we use the following estimator to measure the 3PCF of the galaxies in redshift space:

$$\zeta(s_{12}, s_{23}, s_{31}) = \frac{27RRR^2(s_{12}, s_{23}, s_{31})DDD(s_{12}, s_{23}, s_{31})}{DRR^3(s_{12}, s_{23}, s_{31})} - \frac{9RRR(s_{12}, s_{23}, s_{31})DDR(s_{12}, s_{23}, s_{31})}{DRR^2(s_{12}, s_{23}, s_{31})} + 2.$$
(10)

The above formula is slightly different from the estimator used by Groth & Peebles (1977). Here we have extended the argument of Hamilton (1993) for the 2PCF to the case of the 3PCF. The coefficients 27 and 9 are due to the fact that only *distinct* triplets are counted in this paper. Since the early work of Peebles and coworkers (P80) indicates that the 3PCF of galaxies is approximately hierarchical, it is convenient to express the 3PCF in a normalized form, $Q_{\text{red}}(s_{12}, s_{23}, s_{31})$:

$$Q_{\text{red}}(s_{12}, s_{23}, s_{31}) = \frac{\zeta(s_{12}, s_{23}, s_{31})}{\xi(s_{12})\xi(s_{23}) + \xi(s_{23})\xi(s_{31}) + \xi(s_{31})\xi(s_{12})}.$$
 (11)

It is also convenient to use the variables introduced by Peebles (P80) to describe the shape of the triangles formed by the galaxy triplets. For a triangle with sides $s_{12} \le s_{23} \le s_{31}$, s, u, and v are defined as

$$s = s_{12}, \quad u = \frac{s_{23}}{s_{12}}, \quad v = \frac{s_{31} - s_{23}}{s_{12}}.$$
 (12)

Clearly, u and v characterize the shape and s the size of a triangle. We take equal logarithmic bins for s and u, with the bin intervals $\Delta \log s = \Delta \log u = 0.2$, and equal linear bins for v, with $\Delta v = 0.2$. For our analysis, we take the following ranges for s, u, and v: $0.63 h^{-1} \text{ Mpc} \le s \le 10 h^{-1} \text{ Mpc}$ (6 bins); $1 \le u \le 4$ (3 bins); and $0 \le v \le 1$ (5 bins).

As in JB98, we have generalized the ordinary linked-list technique of P³M simulations (Hockney & Eastwood 1981) to spherical coordinates to count the triplets. The linked-list cells are specified by the spherical coordinates, i.e., the right ascension α , the declination δ , and the distance s. With this

short-range searching technique, we can avoid the triplets out of the range specified, thus making counting triplets very efficient. Because the triplet count RRR is proportional to the third power of the number density of random points, the count within a fixed range of triangles would vary significantly among different luminosity subsamples if the number of random points were fixed, since the volumes covered by different subsamples are very different. We want to have random samples such that the random counts and the cross counts are as big as possible, in order to suppress any uncertainty from the limited number of random points. Therefore, since the CPU time for counting triplets is approximately proportional to the total count of triplets in our linked-list method, we choose a number of random points as large as possible for the computations of RRR, DRR, or RDD, under the condition that each computation is finished in ≤24 CPU hr on a Pentium IV 2.2 GHz PC. The number of random points ranges from 40.000 (for sample IV) to 120,000 (for sample I) when computing RRR and increases to 600,000 (for sample I) when computing *RDD*. The counts *RRR* for small triangles ($s < 1 h^{-1} \text{ Mpc}$) could still be small, and therefore we have recalculated the counts RRR for $s_{31} \le 4 h^{-1}$ Mpc by generating a random sample 10 times larger, so as to ensure that the counts RRR are at least \sim 300 for the triangle configurations of interest. We scaled these counts properly when we determined the 3PCF through equation (10). The uncertainty caused by the number of random points is negligible compared to the sampling errors of the observational sample.

The 3PCF in redshift space $Q_{\rm red}(s,u,v)$ depends on both the real-space distribution of galaxies and their peculiar motions. Although this information contained in $Q_{\rm red}(s,u,v)$ is also useful for the study of the large-scale structures (see § 4), it is apparent that $Q_{\rm red}(s,u,v)$ is different from Q(r,u,v) in real space. By analogy with the analysis for the 2PCF, we have determined the projected 3PCF $\Pi(r_{p12},r_{p23},r_{p31})$. We define the redshift-space 3PCF $\zeta_z(r_{p12},r_{p23},r_{p31},\pi_{12},\pi_{13})$ through

$$dP_{123}^{z} = \bar{n}(s_{1})\bar{n}(s_{2})\bar{n}(s_{3})$$

$$\times \left[1 + \xi_{z}(r_{12}, \, \pi_{12}) + \xi_{z}(r_{23}, \, \pi_{23}) + \xi_{z}(r_{31}, \, \pi_{31}) + \xi_{z}(r_{p_{12}}, \, r_{p_{23}}, \, r_{p_{31}}, \, \pi_{12}, \, \pi_{13})\right] ds_{1} \, ds_{2} \, ds_{3}, \quad (13)$$

where dP_{123}^z is the joint probability of finding one object simultaneously in each of the three volume elements ds_1 , ds_2 , and ds_3 at positions s_1 , s_2 and s_3 , respectively; $\xi_z(r_p, \pi)$ is the redshift space 2PCF; and r_{pij} and π_{ij} are the separations of objects i and j perpendicular to and along the line of sight, respectively. The projected 3PCF $\Pi(r_{p12}, r_{p23}, r_{p31})$ is then defined as

$$\Pi(r_{p12}, r_{p23}, r_{p31}) = \int \zeta_z(r_{p12}, r_{p23}, r_{p31}, \pi_{12}, \pi_{13}) d\pi_{12} d\pi_{23}.$$
(14)

Because the total amount of triplets along the line of sight is not distorted by the peculiar motions, the projected 3PCF $\Pi(r_{p12}, r_{p23}, r_{p31})$ is related to the 3PCF in real space $\zeta(r_{12}, r_{23}, r_{31})$:

$$\Pi(r_{p12}, r_{p23}, r_{p31}) = \int \zeta \left[\sqrt{r_{p12}^2 + y_{12}^2}, \sqrt{r_{p23}^2 + y_{23}^2}, \sqrt{r_{p31}^2 + (y_{12} + y_{23}^2)^2} \right] dy_{12} dy_{23}.$$
(15)

Similarly as for $\zeta(s_{12}, s_{23}, s_{31})$, we measure $\zeta_z(r_{p12}, r_{p23}, r_{p31}, \pi_{12}, \pi_{13})$ by counting the numbers of triplets $DDD(r_{p12}, r_{p23}, r_{p31}, \pi_{12}, \pi_{13})$, $DRR(r_{p12}, r_{p23}, r_{p31}, \pi_{12}, \pi_{13})$, $RDD(r_{p12}, r_{p23}, r_{p31}, \pi_{12}, \pi_{13})$, and $RRR(r_{p12}, r_{p23}, r_{p31}, \pi_{12}, \pi_{13})$ formed by galaxies and/or random points with the projected separations r_{p12}, r_{p23} , and r_{p31} and radial separations π_{12} and π_{23} . We use r_p, u , and v,

$$r_p = r_{p12}, \quad u = \frac{r_{p23}}{r_{p12}}, \quad v = \frac{r_{p31} - r_{p23}}{r_{p12}},$$
 (16)

to quantify a triangle with $r_{p12} \le r_{p23} \le r_{p31}$ on the plane perpendicular to the line of sight. Equal logarithmic bins of intervals $\Delta \log r_p = \Delta \log u = 0.2$ are taken for r_p and u and equal linear bins of $\Delta v = 0.2$ for v. The same ranges of u and v are used as for $\zeta(s, u, v)$, but r_p is from 0.128 to 4 h^{-1} Mpc (7 bins). The radial separations π_{12} and π_{23} are from -25 to $25 \ h^{-1}$ Mpc, with a bin size of $1 \ h^{-1}$ Mpc. The projected 3PCF is estimated by summing up $\zeta_z(r_p, u, v, \pi_{12}^i, \pi_{23}^j)$ at different radial bins (π_{12}^i, π_{23}^j) ,

$$\Pi(r_p, u, v) = \sum_{i,j} \zeta_z(r_p, u, v, \pi_{12}^i, \pi_{23}^j) \Delta \pi_{12}^i \Delta \pi_{23}^j, \quad (17)$$

and it is normalized as

$$Q_{\text{proj}}(r_{p}, u, v) = \frac{\Pi(r_{p}, u, v)}{w(r_{p12})w(r_{p23}) + w(r_{p23})w(r_{p31}) + w(r_{p31})w(r_{p12})},$$
(18)

where $w(r_p)$ is the projected 2PCF (Davis & Peebles 1983; JMB98)

$$w(r_p) = \sum_i \xi_z(r_p, \pi^i) \Delta \pi^i.$$
 (19)

An interesting property of the projected 3PCF is that if the 3PCF is of the hierarchical form, the normalized function $Q_{\text{proj}}(r_p, u, v)$ is not only a constant but also equal to Q. Therefore, the measurement of $Q_{\text{proj}}(r_p, u, v)$ can be used to test the hierarchical form, which was proposed mainly based on the analysis of angular catalogs.

JB98 have used *N*-body simulations to test the statistical methods for the LCRS and found that the results obtained are unbiased. Since the 2dFGRS is constructed in a similar way to the LCRS and the survey area is larger, the above methods should also yield unbiased results for the 2dFGRS.

The error bars of Q are estimated by the bootstrap method (Barrow, Bhavsar, & Sonoda 1984; Mo, Jing, & Börner 1992). We have also used the mock samples of DM particles in \S 4 to estimate the error bars. We find that the error bars from these two methods agree within a factor of 2. Here we adopt the bootstrap error for the measurement of Q, since we do not input a luminosity-dependent bias for mock samples.

As in the analysis of the 2PCF, the estimates of the 3PCF given by equation (10) are correlated on different scales. This point should be taken into account when the measured 3PCF is compared with model predictions. Recently, new techniques have been developed to tackle this important issue in the context of the 2PCF or the power spectrum, e.g., Tegmark, Hamilton, & Xu (2002) and Matsubara & Szalay (2002), using

the Karhunen-Loève eigenmode analysis, and Fang & Feng (2000) and Zhan, Jamkhedkar, & Fang (2001), using the wavelet analysis. It remains an important task to study if these methods can be extended to obtain a decorrelated 3PCF.

4. RESULTS OF THE 2dFGRS ANALYSIS

4.1. The 3PCF of the 2dFGRS Catalog and the Luminosity Dependence

We present the results for the 3PCF in redshift space $Q_{\rm red}(s,u,v)$ and for the projected 3PCF $Q_{\rm proj}(r_p,u,v)$ in Figures 3–12 for the 2dFGRS. The errors of the Q-values are estimated by the bootstrap resampling method. The large number of galaxies in the 2dFGRS allows us to look for a possible luminosity dependence of $Q_{\rm red}(s,u,v)$ and $Q_{\rm proj}(r_p,u,v)$. We have selected five galaxy samples according to luminosity, listed in Table 1. The samples are not completely independent, with significant overlaps between some of the samples.

For $Q_{\rm red}(s,u,v)$ the results are shown in Figures 3–7. As we can see, the 3PCF in redshift space does not change very much with s or u and increases somewhat with v for fixed s and u. For small v $Q_{\rm red}(s,u,v)$ is approximately constant at a value of \sim 0.6, but it increases to \sim 1 when v \sim 1.

For the bright galaxies we find that $Q_{\rm red}(s,u,v)$ decreases somewhat with s, from 0.9 at $s=0.82~h^{-1}$ Mpc to 0.4 at $s=5.15~h^{-1}$ Mpc. Changes with s are slightly reduced for the samples including fainter galaxies. For the faintest sample (IV), at small s and v $Q_{\rm red}(s,u,v)$ is about 1.1, and it decreases to \sim 0.7 at $s=3.25~h^{-1}$ Mpc.

We find that $Q_{\text{red}}(s, u, \bar{v})$ is slightly larger for the fainter samples, although the dependence on luminosity is rather weak. In fact, if the errors are taken into account, this luminosity dependence is not statistically significant. We also note that there is always some difference between the NGP strip, the SGP strip, and the whole sample, but generally within the 1 σ error bars. This implies that the bootstrap error used in this sample is a good indicator for the error estimate. The results for the north and the south samples are in good agreement for the galaxies brighter than $M_b - 5 \log h \le -18.5$. For the faint sample with $M_h - 5 \log h > -18.5$, however, there is a significant difference between the north and south subsamples. The main reason is that this sample covers only a small cosmic volume, so the sample-to-sample difference (the cosmic variance) can be large. In fact, even the 2PCFs of these subsamples are dramatically different (see Fig. 1). Considering the fact that the bootstrap error is not sufficient to fully account for the cosmic variance, one should remain cautious about the results for the faintest sample (IV). Nonetheless, from Figures 3-7 we conclude that there is at best a slight dependence on luminosity, in the sense that the amplitude $Q_{\text{red}}(s, u, v)$ tends to be smaller for brighter galaxies.

The projected 3PCF, in comparison, shows a behavior that is somewhat different. In the bright galaxy sample (Fig. 8), $Q_{\text{proj}}(r_p, u, v)$ is about 0.7 at $r_p = 0.2 \ h^{-1}$ Mpc, and it reaches down to $Q_{\text{proj}}(r_p, u, v) \simeq 0.5$ at $r_p = 3.25 \ h^{-1}$ Mpc for small v, so the dependence on r_p is quite mild. There is, however, a small but significant increase with v. Fainter galaxies show a similar weak dependence on r_p and v (Figs. 9 and 11). Comparing different samples, however, we find a trend that brighter galaxies have lower values of $Q_{\text{proj}}(r_p, u, v)$. The $Q_{\text{proj}}(r_p, u, v)$ of the fainter samples (II and IV) is about 50% higher than that of the brightest sample,

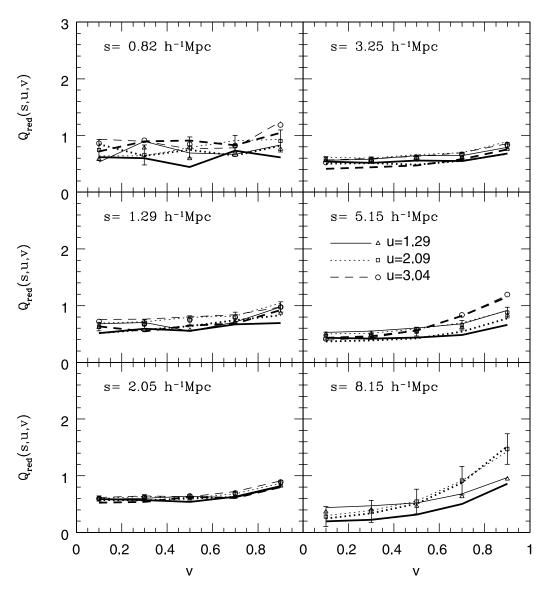


Fig. 3.—Normalized 3PCF in redshift space $Q_{\rm red}(s,u,v)$ of galaxies with luminosity $M_b-5\log h \le -19.66$ in the 2dFGRS. The results for the SGP strip, the NGP strip, and the whole sample are plotted with thick lines, thin lines, and symbols, respectively. Different lines and symbols are used for triangle configurations of different u-values, as indicated. The errors are estimated by the bootstrap resampling method. For clarity, the error bars are plotted for the whole sample and u=2 only, but those for the other two u-values are very similar, and those for the NGP and SGP strips are about 1.4 times larger.

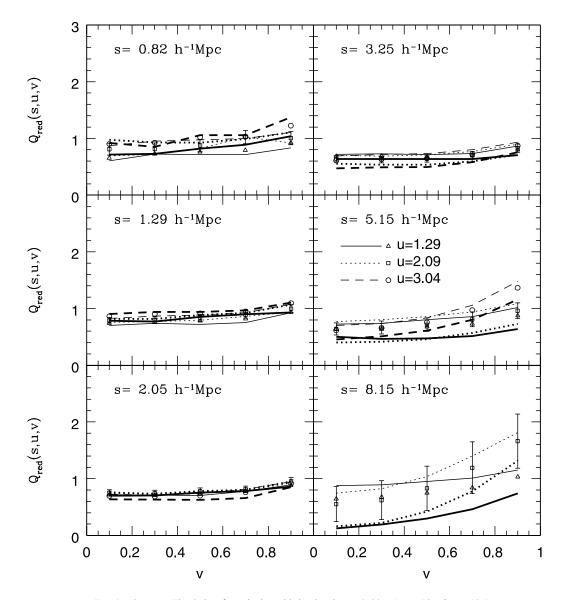


Fig. 4.—Same as Fig. 3, but for galaxies with luminosity $-19.66 < M_b - 5 \log h \leq -18.5$

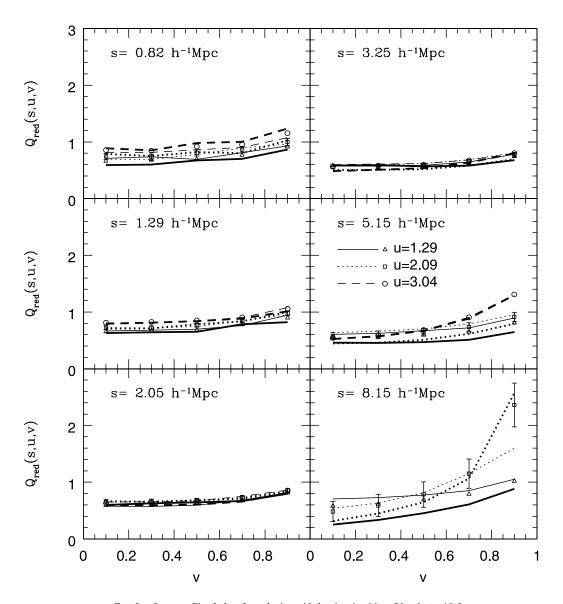


Fig. 5.—Same as Fig. 3, but for galaxies with luminosity $M_b - 5 \log h \le -18.5$

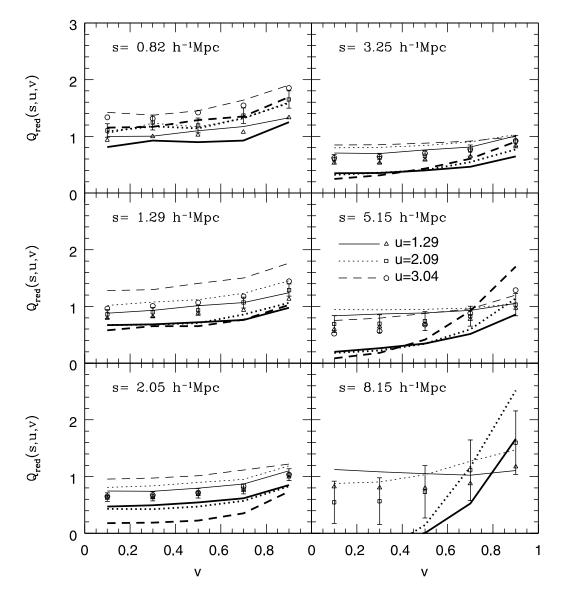


Fig. 6.—Same as Fig. 3, but for galaxies with luminosity $M_b - 5\log h > -18.5$

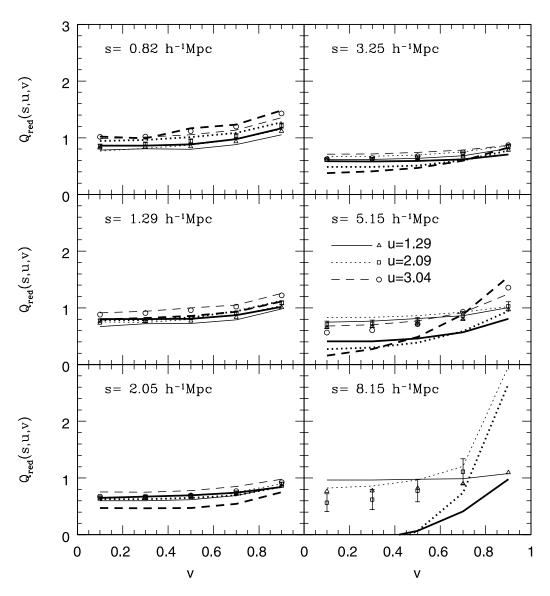


Fig. 7.—Same as Fig. 3, but for all galaxies in the 2dFGRS (without luminosity selection)

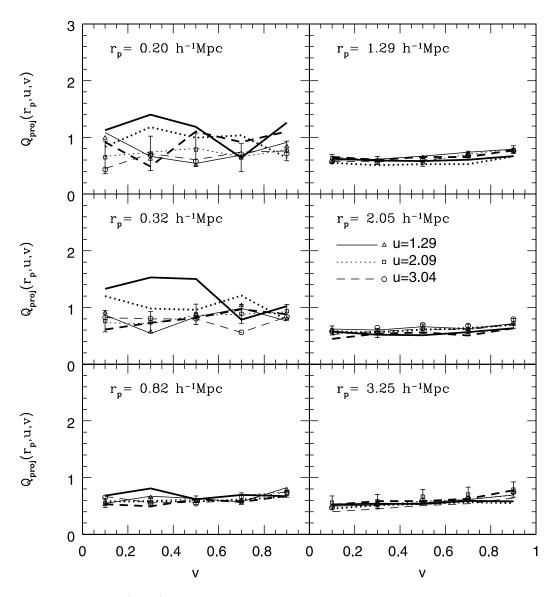


Fig. 8.—Normalized projected 3PCF $Q_{\text{proj}}(r_p, u, v)$ of galaxies with luminosity $M_b - 5 \log h \le -19.66$ in the 2dFGRS. Lines, symbols, and error bars as in Fig. 3.

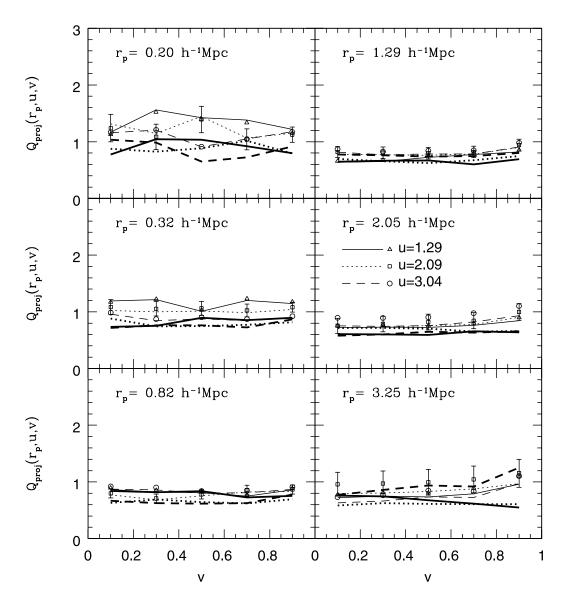


Fig. 9.—Same as Fig. 8, but for galaxies with luminosity $-19.66 < M_b - 5 \log h \le -18.5$

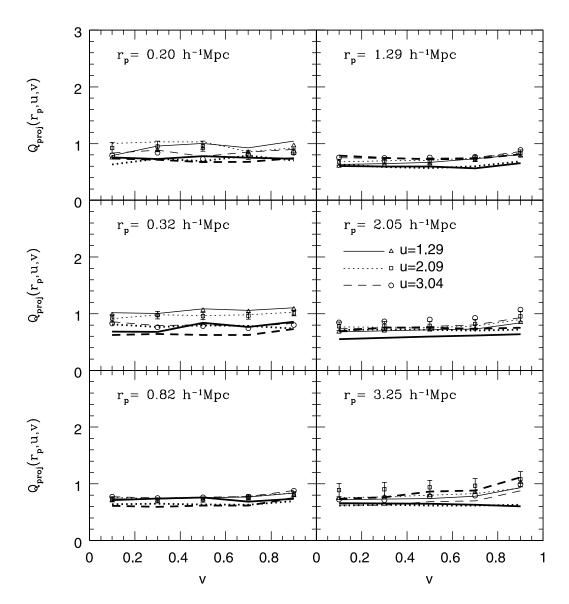


Fig. 10.—Same as Fig. 8, but for galaxies with luminosity $M_b-5\log h \leq -18.5$

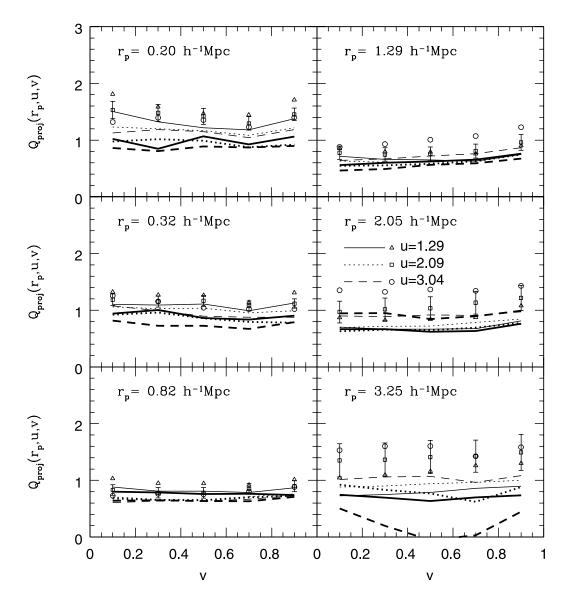


Fig. 11.—Same as Fig. 8, but for galaxies with luminosity $M_b - 5 \log h > -18.5$

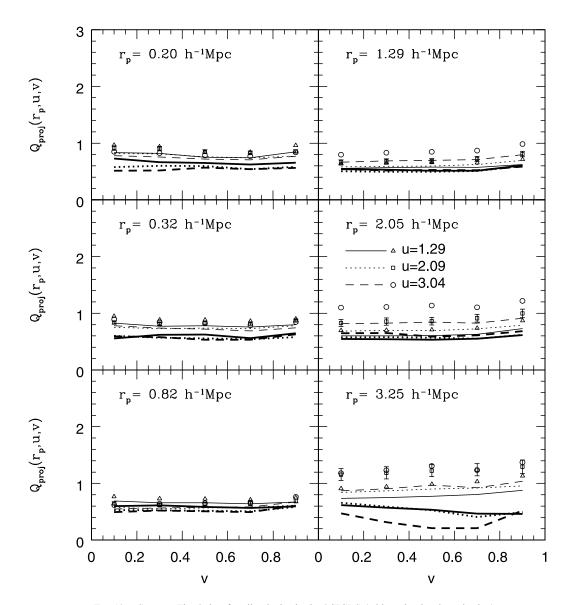


Fig. 12.—Same as Fig. 8, but for all galaxies in the 2dFGRS (without luminosity selection)

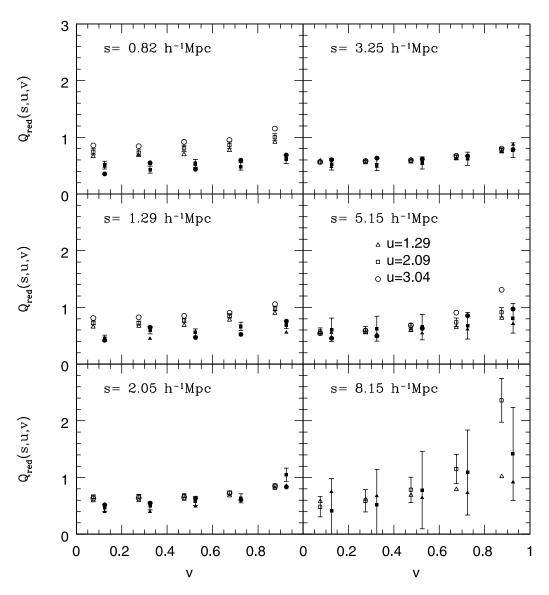


Fig. 13.—Comparison of the normalized 3PCFs in redshift space $Q_{\rm red}(s,u,v)$ measured from the 2dFGRS and from the LCRS. The LCRS data are taken from JB98 for galaxies with R-band luminosity $M_R-5\log h \le -18.5$. For comparison, we simply take our result in the 2dFGRS of galaxies with $M_b-5\log h \le -18.5$, despite the fact that galaxies are selected in different wave bands in the two surveys. The LCRS results are plotted in solid symbols, and 2dFGRS results are in open symbols. The errors are estimated by the bootstrap resampling method and are plotted for u=2 only. For clarity, the symbols are shifted by +0.05 for the LCRS and by -0.05 for 2dFGRS along the horizontal axis.

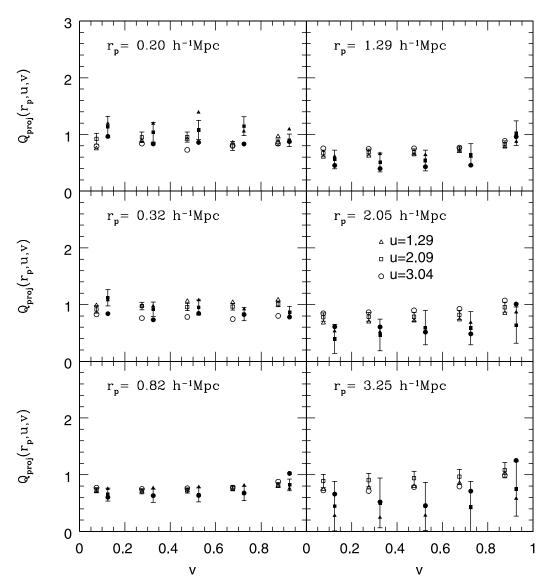


Fig. 14.—Same as Fig. 13, but for the normalized projected 3PCFs $Q_{\text{proj}}(r_p, u, v)$ from the 2dFGRS and the LCRS

 $M_b - 5 \log h < -19.66$ (I). We discuss the implications for the bias parameters in $\S 4.3$.

Figure 11 shows that while the values of $Q_{\text{proj}}(r_p, u, v)$ are similar for the north and south subsamples, the value for the total sample is larger than that for either subsample. This looks a bit surprising at first glance. But considering that the 2PCF of the north sample is almost 1.5 times larger than that of the south sample on h^{-1} Mpc scales, it is not difficult to explain the behavior of $Q_{\text{proj}}(r_p, u, v)$ for the total sample and the two subsamples. As an idealized example, we assume that the two subsamples are well separated and have the same sample size and $Q_{\text{proj}}(r_p, u, v)$, but the 2PCF of one sample is 1.5 times larger than that of the other. This example is quite close to the real situation of the faintest sample. It is not difficult to prove that the $Q_{\text{proj}}(r_p, u, v)$ of the total sample is 1.4 times that of the subsamples. With this example, it is easy to see that the amplitude $Q_{\text{proj}}(r_p, u, v)$ of the total sample is larger than that of the subsamples for $r_p \sim 1.29$ to $\sim 3.25 \ h^{-1}$ Mpc for the faintest galaxies. This unusual behavior again can be attributed to the fact that this sample surveys only a small volume of sky and that the cosmic variance is therefore large.

4.2. Comparison with Results from the LCRS

In Figure 13, we compare the normalized 3PCFs in redshift space, $Q_{\text{red}}(s, u, v)$, of the 2dFGRS and LCRS. The LCRS data are taken from JB98 for galaxies with luminosities in the R band $M_R - 5 \log h \le -18.5$. From the 2dFGRS we simply take our result for the galaxies with $M_b - 5 \log h \le -18.5$, although we are aware of the fact that the galaxies are selected in different wave bands in the two surveys. There are subtle differences in the results, which we attribute to this choice of the observational bands, because Q depends on luminosity weakly for $M > M^*$. For small values of $s \sim 1$ h^{-1} Mpc, the 2dF catalog gives a slightly higher amplitude than the LCRS galaxies. This could reflect the fact that the real-space 2PCF of the LCRS galaxies is higher than that of the APM galaxies on small scales, as JMB98 pointed out. Nevertheless, the $Q_{\text{red}}(s, u, v)$ values agree very well between the two samples, especially on larger scales. The 2dF sample gives rise to a much smaller error, because of its large sample size.

To compare the projected amplitudes $Q_{\text{proj}}(r_p, u, v)$, we display this quantity for the two catalogs in Figure 14. Again

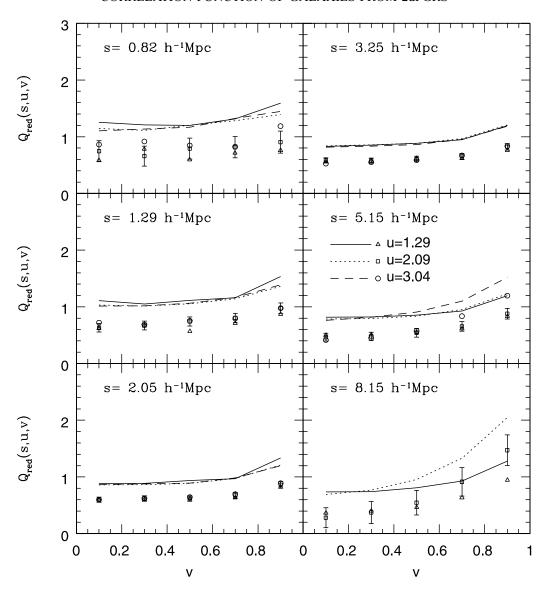


Fig. 15.—Comparison of the normalized 3PCF of galaxies in redshift space $Q_{\text{red}}(s, u, v)$ with the function predicted in the WMAP running power CDM model. The observed data (symbols) are from the 2dFGRS for $M_b - 5 \log h \le -19.66$, and the model prediction is for DM (lines). The errors are estimated by the bootstrap resampling method and are plotted for u = 2 only.

the agreement is quite satisfactory, especially when we take the larger error bars for the LCRS result into consideration. However, the systematic decrease with r_p that can be read off for the mean values of $Q_{\rm proj}(r_p,u,v)$ for the LCRS data is not present for the 2dFGRS. This is probably caused by the fact that the sky area of the LCRS is much smaller than that of the 2dFGRS; the mean value for the LCRS is thus systematically underestimated. The 2dFGRS data also imply that the real-space 3PCF of galaxies on the small scales explored here does not deviate significantly from the hierarchical form (P80) and that the fitting formula given in JB98 for the projected $Q_{\rm proj}(r_p,u,v)$ needs to be revised.

In conclusion, our 2dFGRS results for Q, in both redshift and projected space, are in good agreement with the results obtained by JB98 for the LCRS.

4.3. Comparison with the DM Distribution in the Running Power CDM Model

In this section, we compare the observational results with model predictions. Currently, the parameters of the CDM model have been determined pretty accurately by a combination of data from WMAP (Wilkinson Microwave Anisotropy *Probe*), 2dFGRS, and Ly α absorption systems, and complementarily by many other observations (Spergel et al. 2003). We choose the running power CDM model of Spergel et al. for comparison with our statistical results, for this model matches most available observations: the universe is flat, with a density parameter $\Omega_0 = 0.26$ and a cosmological constant $\lambda_0 = 0.74$. The Hubble constant is 71 km s⁻¹ Mpc⁻¹, and the baryonic density parameter $\Omega_{0,b} = 0.045$. The primordial density power spectrum deviates slightly from the Zel'dovich spectrum as $P(k) \propto (k \text{ Mpc}/0.05)^{n(k)}$, with n(k) = 0.93 - 0.0165 ln(k Mpc/0.05). Although there is no consensus about the necessity of introducing the running power index n(k) (e.g., Seljak, McDonald, & Makarov 2003; Tegmark et al. 2004), we choose this model as a reasonable approximation to the

Because the 3PCFs that we have measured are in the non-linear and quasi-linear regimes, we use an *N*-body simulation to make model predictions. The simulation has 512³ particles

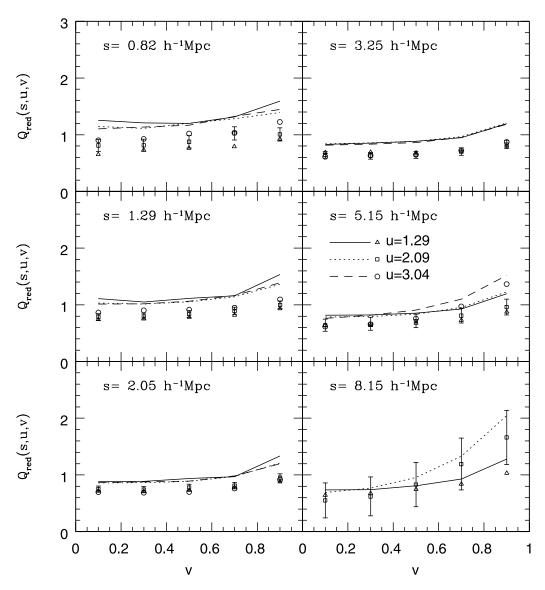


Fig. 16.—Same as Fig. 15, but the observed data (symbols) are from the 2dFGRS for $-19.66 < M_b - 5 \log h \le -18.5$

in a cubic box of $1024\ h^{-1}$ Mpc and is generated with our P^3M code (see Jing & Suto 2002 for the code). To include the effect of baryonic matter oscillations on large-scale structures, the fitting formula of Eisenstein & Hu (1999) for the transfer function is used to generate the initial conditions. Since the median redshift of the 2dFGRS is ~ 0.13 , we choose the simulation output at this redshift. We note that the three-point correlation is quite sensitive to the presence of very massive clusters; therefore, a large simulation box like the one used here is necessary. With a small box of $\lesssim 100\ h^{-1}$ Mpc, the 3PCF may be underestimated severely.

Generally speaking, galaxies are biased tracers of the underlying matter distribution in the universe. A luminosity dependence of the bias (Norberg et al. 2002a) means that faint and bright galaxies trace the matter distribution differently. It has become popular in recent years to account for the bias of certain types of galaxies phenomenologically with the so-called halo occupation number model (e.g., JMB98; Seljak 2000; Peacock & Smith 2000; Sheth et al. 2001; Berlind & Weinberg 2002; Cooray & Sheth 2002; Zehavi et al. 2004; see Yang et al. 2003 for an updated account of this model). The 3PCF of galaxies can also be modeled within this framework

(JB98; Berlind & Weinberg 2002; Ma & Fry 2000; see Takada & Jain 2003 for a detailed account of this modeling), although it seems difficult to account for the 2PCFs and 3PCFs in the LCRS simultaneously with simple power-law occupation models (JB98). Our accurate measurement of the 3PCF for the 2dFGRS and its luminosity-dependence will certainly provide an even more stringent constraint on the halo occupation number models. It remains to be seen if the sophisticated model of Yang et al. (2003, 2004) can explain the results obtained in this paper. We want to investigate this issue in a subsequent paper, and here we only compare with one model prediction for the DM, in order to set a baseline quantifying the difference in the normalized 3PCF between real galaxies and DM for the concordance CDM model.

The comparison between the 2dFGRS results and the model predictions is displayed in Figures 15–18. Here we consider only two luminosity subsamples. First, we find that the qualitative features, such as the dependence on v for fixed $s(r_p)$ and u, the decrease of Q with increasing values of s or r_p are reproduced quite well by the DM simulations. For the luminous sample (sample I), the Q-values of the data set are generally lower than the DM model predictions, up to a factor

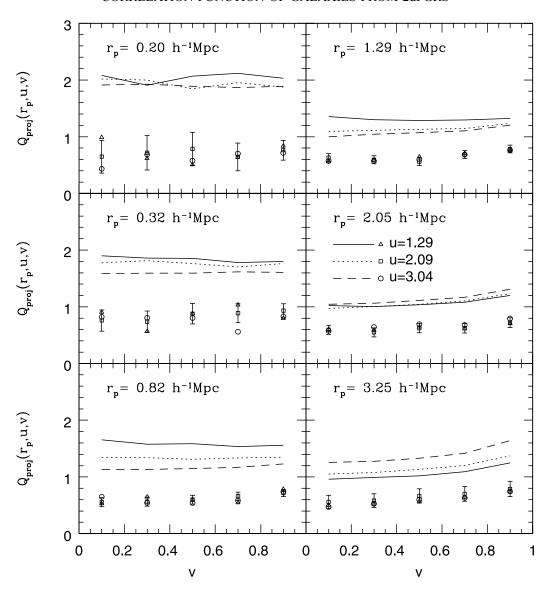


Fig. 17.—Comparison of the normalized projected 3PCF of galaxies $Q_{\text{proj}}(r_p, u, v)$ with the function predicted in the WMAP running power CDM model. The observed data (symbols) are from the 2dFGRS for $M_b - 5 \log h \le -19.66$. Lines and errors as in Fig. 15.

of \sim 1.5 to \sim 2. For the less luminous sample (sample II), the observed Q-values also are smaller than those for the DM model on small scales, but the observed values and the model predictions agree at the values $s \approx 8 \ h^{-1}$ Mpc and $r_p = 3.25 \ h^{-1}$ Mpc. Because the largest scales probed here are expected to be linear or quasi-linear scales, we expect the linear bias model (eq. [2]) to hold on these scales. Our result therefore tells us that on linear scales, the galaxies with $-19.66 < M_b - 5 \log h \le -18.5$ are approximately an unbiased tracer, but the brightest galaxies, with $M_b - 5 \log h \le -19.66$, have a bias factor of \sim 1.5.

Because the 2PCF of the galaxies of sample II matches well the 2PCF of the DM in the concordance WMAP model and our 3PCF results show that the galaxies of sample II are unbiased on large, linear scales, we find support for the density fluctuation normalization $\sigma_8 = 0.84$ obtained by Spergel et al. (2003). On the other hand, our result shows that the three-point correlations Q of galaxies are lower on nonlinear scales than the prediction of the WMAP concordance model. Physical models, e.g., the halo occupation number model (e.g., Yang

et al. 2003) or the semianalytical models of galaxy formation (e.g., Kauffmann et al. 1997), are needed to interpret the observed small-scale, nonlinear bias. We will pursue this in a future paper. The three-point correlation amplitudes of samples III and V are very close to that of sample II. The Q of these samples gradually conforms to the model prediction of the concordance model on quasi-linear scales $r \sim 5 \ h^{-1}$ Mpc. Our results are therefore consistent with the analysis of Verde et al. (2002), who showed that the 2dFGRS galaxies (without a luminosity classification) are an unbiased tracer of the underlying matter on scales of 5–30 h^{-1} Mpc.

5. CONCLUSION

In a detailed analysis of the 3PCF for the 2dFGRS, we have accurately measured the 3PCF for galaxies of different luminosities. The 3PCF amplitudes $[Q_{\rm red}(s,u,v) \text{ or } Q_{\rm proj}(r_p,u,v)]$ of galaxies generally decrease with the increase of the triangle size and increase with the increase of v, qualitatively in agreement with the predictions for the DM clustering in popular hierarchical CDM models. Some dependence on

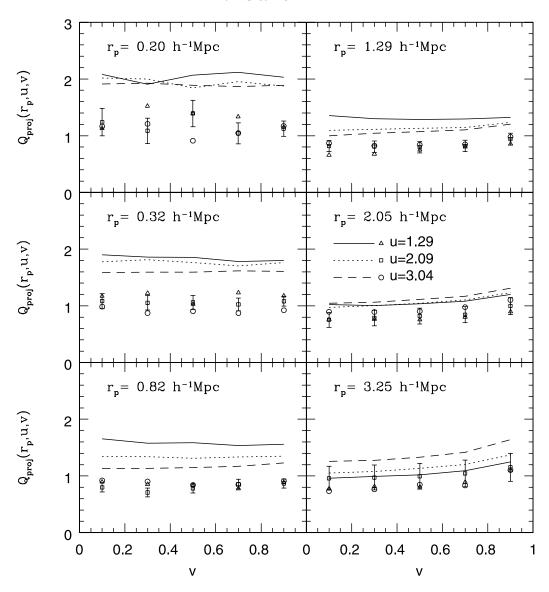


Fig. 18.—Same as Fig. 17, but the observed data (symbols) are from the 2dFGRS for $-19.66 < M_b - 5 \log h \le -18.5$

luminosity is found, but not a strong effect, except for the brightest galaxy sample, which seems to have amplitudes lower by up to 50%. Comparing with the previous study on the LCRS galaxies (JB98), we find good agreement between the two studies, although the results from the 2dFGRS are more accurate, since the 2dFGRS is much larger than the LCRS. The amplitudes in redshift space $Q_{\rm red}(s,u,v)$ are very similar, but the projected ones $Q_{\rm proj}(r_p,u,v)$ show some difference. It seems that the projected 3PCF from the LCRS is systematically underestimated for r_p in the range of a few h^{-1} Mpc, because of the thin-slice geometry of that survey. The dependence of $Q_{\rm proj}(r_p,u,v)$ on r_p is much weaker in the 2dFGRS than in the LCRS.

Comparing the measured 3PCF amplitudes $[Q_{\rm red}(s,u,v)]$ or $Q_{\rm proj}(r_p,u,v)]$ to the predictions of a *WMAP* concordance model, we find that the measured values are consistently lower than the predicted ones for DM. As in the case of the LCRS, about one-half of the predicted Q-value provides a good description of $Q_{\rm proj}(r_p,u,v)$ for the 2dFGRS data. As in JB98 for the LCRS data, we may state that the best-fit DM model gives higher values for the 3PCF than observed. This indicates that

the simple DM models must be refined, either by using more sophisticated bias models or by a more sophisticated combination of model parameters.

The division of galaxies into luminosity classes reveals that the brightest galaxies are biased even on large scales, while the galaxies of sample II (Table 1) show a nonlinear bias on small scales but appear unbiased on linear scales.

This work has made use of the data released by the 2dFGRS team and the software for generating mock samples provided by Peder Norberg and Shaun Cole. We are grateful to Peder Norberg for his explanations on how to use the masks and the software. We are grateful to Volker Müller, Bob Nichol, and Yasushi Suto for communicating their SDSS results of the three-point correlations before publication. Y. P. J. would like to thank the Max-Planck-Institut für Astrophysik for its warm hospitality during the time the work was completed. The work is supported in part by NKBRSF (G19990754), by NSFC (10125314), and by the CAS-MPG exchange program.

REFERENCES

Bahcall, N. A., & Comerford, J. M. 2002, ApJ, 565, L5

Barriga, J., & Gaztañaga, E. 2002, MNRAS, 333, 443

Barrow, J. D., Bhavsar, S. P., & Sonoda, D. H. 1984, MNRAS, 210, 19P

Bartelmann, M., & Schneider, P. 2001, Phys. Rep., 340, 291

Bean, A. J., Efstathiou, G., Ellis, R. S., Peterson, B. A., & Shanks, T. 1983, MNRAS, 205, 605

Benson, A. J., Baugh, C. M., Cole, S., Frenk, C. S., & Lacey, C. G. 2000, MNRAS, 316, 107

Berlind, A. A., & Weinberg, D. H. 2002, ApJ, 575, 587

Bernardeau, F., Colombi, S., Gaztañaga, E., & Scoccimarro, R. 2002, Phys. Rep., 367, 1

Börner, G., Deng, Z.-G., Xia, X.-Y., & Zhou, Y.-Y. 1991, Ap&SS, 180, 47 Buchalter, A., & Kamionkowski, M. 1999, ApJ, 521, 1

Catelan, P., Lucchin, F., Matarrese, S., & Porciani, C. 1998, MNRAS, 297, 692

Colless, M., et al. 2001, MNRAS, 328, 1039 (C01) Cooray, A., & Sheth, R. 2002, Phys. Rep., 372, 1

Davis, M., Efstathiou, G., Frenk, C. S., & White, S. D. M. 1985, ApJ, 292, 371

Davis, M., & Peebles, P. J. E. 1983, ApJ, 267, 465

Efstathiou, G., & Jedrzejewski, R. I. 1984, Adv. Space Res., 3, 379

Eisenstein, D. J., & Hu, W. 1999, ApJ, 511, 5

Fang, L.-Z., & Feng, L. 2000, ApJ, 539, 5

Frieman, J. A., & Gaztañaga, E. 1999, ApJ, 521, L83

Fry, J. N. 1984, ApJ, 279, 499

Gaztañaga, E., & Frieman, J. A. 1994, ApJ, 437, L13

Groth, E. J., & Peebles, P. J. E. 1977, ApJ, 217, 385

Hale-Sutton, D., Fong, R., Metcalfe, N., & Shanks, T. 1989, MNRAS, 237, 569

Hamilton, A. J. S. 1993, ApJ, 417, 19

Hamilton, A. J. S., & Tegmark, M. 2002, MNRAS, 330, 506

Hawkins, E., et al. 2003, MNRAS, 346, 78

Hockney, R. W., & Eastwood, J. W. 1981, Computer Simulations Using Particles (New York: McGraw-Hill)

Jing, Y. P., & Börner, G. 1997, A&A, 318, 667

1998, ApJ, 503, 37 (JB98)

Jing, Y. P., Mo, H. J., & Börner, G. 1991, A&A, 252, 449

1998, ApJ, 494, 1 (JMB98)

Jing, Y. P., & Suto, Y. 2002, ApJ, 574, 538

Kauffmann, G., Colberg, J. M., Diaferio, A., & White, S. D. M. 1999, MNRAS, 303, 188

Kauffmann, G., Nusser, A., & Steinmetz, M. 1997, MNRAS, 286, 795 Lahav, O., et al. 2002, MNRAS, 333, 961

Lauberts, A. 1982, The ESO-Uppsala Survey of the ESO(B) Atlas (Garching: ESO)

Loveday, J., Maddox, S. J., Efstathiou, G., & Peterson, B. A. 1995, ApJ,

Ma, C.-P., & Fry, J. N. 2000, ApJ, 531, L87

Matarrese, S., Verde, L., & Heavens, A. F. 1997, MNRAS, 290, 651

Matsubara, T. 1994, ApJ, 424, 30

Matsubara, T., & Szalay, A. S. 2002, ApJ, 574, 1

Mo, H. J., Jing, Y. P., & Börner, G. 1992, ApJ, 392, 452

Mo, H. J., Jing, Y. P., & White, S. D. M. 1997, MNRAS, 284, 189

Norberg, P., et al. 2002a, MNRAS, 332, 827

2002b, MNRAS, 336, 907

Peacock, J. A., & Smith, R. E. 2000, MNRAS, 318, 1144

Peebles P. J. E. 1980, The Large-Scale Structure of the Universe (Princeton: Princeton Univ. Press) (P80)

1981, in Ann. NY Acad. Sci., 375, Proc. 10th Texas Symp. on Relativistic Astrophysics, ed. R. Ramaty & F. C. Jones, 157

Pen, U.-L., Lu, T., van Waerbeke, L., & Mellier, Y. 2003, MNRAS, 346, 994 Scoccimarro, R., Colombi, S., Fry, J. N., Frieman, J. A., Hivon, E., & Melott, A. 1998, ApJ, 496, 586

Scoccimarro, R., Feldman, H. A., Fry, J. N., & Frieman, J. A. 2001, ApJ, 546, 652

Seljak, U. 2000, MNRAS, 318, 203

Seljak, U., McDonald, P., & Makarov, A. 2003, MNRAS, 342, L79

Shectman, S. A., Landy, S. D., Oemler, A., Tucker, D. L., Lin, H., Kirshner, R. P., & Schechter, P. L. 1996, ApJ, 470, 172

Sheth, R. K., Hui, L., Diaferio, A., & Scoccimarro, R. 2001, MNRAS, 325, 1288

Spergel, D. N., et al. 2003, ApJS, 148, 175

Takada, M., & Jain, B. 2003, MNRAS, 340, 580

Tegmark, M., Hamilton, A. J. S., & Xu, Y. 2002, MNRAS, 335, 887

Tegmark, M., et al. 2004, Phys. Rev. D, in press

van den Bosch, F. C., Mo, H. J., & Yang, X. 2003, MNRAS, 345, 923 Verde, L., et al. 2002, MNRAS, 335, 432

Xia, X. Y., Deng, Z. G., & Zhou, Y. Y. 1987, in IAU Symp. 124, Observational Cosmology, ed. A. Hewitt, G. Burbridge, & L.-Z. Fang (Dordrecht: Reidel),

Yang, X., Mo, H. J., Jing, Y. P., van den Bosch, F. C., & Chu, Y.-Q. 2004, MNRAS, in press

Yang, X., Mo, H. J., & van den Bosch, F. C. 2003, MNRAS, 339, 1057

Zehavi, I., et al. 2002, ApJ, 571, 172

2004, ApJ, in press

Zhan, H., Jamkhedkar, P., & Fang, L. 2001, ApJ, 555, 58